

Math 180A: Introduction to Probability

Lecture B00 (Nemish)

math.ucsd.edu/~ynemish/teaching/180a

Lecture C00 (Au)

math.ucsd.edu/~bau/f20.180a

Today: ASV 6.1, 6.2

Video: Prof. Todd Kemp, Fall 2019

Next: ASV 6.2, 6.3

Week 8: Homework 7 (due Friday, Dec 4)

Regrades for Homework 5 (Nov 30 - Dec 1)

Definition: Given (discrete) random variables X_1, X_2, \dots, X_n all defined on the same sample space, their joint distribution is the collection of all

$$\left\{ \begin{array}{l} \mathbb{P}(X_1=k_1, X_2=k_2, \dots, X_n=k_n) \\ \text{all possible values } k_1 \text{ of } X_1, k_2 \text{ of } X_2, \dots, k_n \text{ of } X_n \end{array} \right\}$$

Eg. $X, Y \sim \text{Ber}(p)$,

(1) $X = Y$

(2) X, Y independent

$$\mathbb{P}(X=k_1, Y=k_2)$$

$$= \mathbb{P}(X=k_1) \mathbb{P}(Y=k_2)$$

Y {

0

1

		X	
		0	1
0	1-p	$(1-p)^2$	0
	0	$p(1-p)$	$p(1-p)$
1	0	$(1-p)p$	0
	1	p	p^2

$$P_{X,Y}(k_1, k_2) = \mathbb{P}(X=k_1, Y=k_2)$$

Recovering X_j from $\underline{X} = (X_1, X_2, \dots, X_n)$: Marginals

Suppose we know $P_{\underline{X}}(\underline{k})$ for all $\underline{k} = (k_1, k_2, \dots, k_n)$.

How can we find $P_{X_1}(t)$?

Eg. Toss a fair coin twice. $X_1, X_2 \in \{0, 1\}$ $\frac{1}{4} + \frac{1}{4} = \frac{1}{2}$.

$$P(X_1=0) = P(X_1=0, X_2=0) + P(X_1=0, X_2=1)$$

$$\{X_1=0\} = \{X_1=0, X_2=0\} \cup \{X_1=0, X_2=1\}$$

In general,

$$P(X_1=t) = \sum_{k_2, k_3, \dots, k_n} P(X_1=t, X_2=k_2, \dots, X_n=k_n) = \sum_{k_2, \dots, k_n} P_{\underline{X}}(t, k_2, k_3, \dots, k_n)$$

$$\{X_1=t\} = \bigcup_{k_2, \dots, k_n} \{X_1=t, X_2=k_2, X_3=k_3, \dots, X_n=k_n\}$$

$$P(X=t) = \sum_k P_{(X,Y)}(t, k)$$

(X, Y)

$$P(Y=t) = \sum_k P_{(X,Y)}(k, t)$$

Eg. Toss a fair coin 3 times. $X = \# \text{tails in first toss (0 or 1)}$
 $Y = \text{total \#tails in all 3 (0, 1, 2, 3)}$

		Y							
		0	1	2	3		Outcome	X	Y
X	0	$\frac{1}{8}$	$\frac{1}{4}$	$\frac{1}{8}$	0	$\rightarrow \frac{1}{2}$	HHH	0	0
	1	0	$\frac{1}{8}$	$\frac{1}{4}$	$\frac{1}{8}$	$\rightarrow \frac{1}{2}$	HHT	0	1
		\downarrow	\downarrow	\downarrow	\downarrow	$\leftarrow B_m(3, \frac{1}{2})$	HTH	0	1
		$\frac{1}{8}$	$\frac{3}{8}$	$\frac{3}{8}$	$\frac{1}{8}$		HTT	0	2
							TTH	1	1
							THT	1	2
							TTH	1	2
							TTT	1	3

each has $P = \frac{1}{8}$.

$P(X=k, Y=l)$
 0 0
 1 1
 2 2
 3 3

Question: Are X, Y independent?

$P(X=1, Y=0) = 0$

No!

$P(X=1)P(Y=0) = \frac{1}{2} \cdot \frac{1}{8} = \frac{1}{16} \neq 0$.

Joint distributions are just distributions of random vectors.

$$X_1, X_2, \dots, X_n \rightsquigarrow \underline{X} = (X_1, X_2, \dots, X_n)$$

possible values for \underline{X} are vectors (k_1, \dots, k_n) .

Eg. Multinomial Distribution.

Often, trials have more than 2 outcomes.

Consider a trial w r possible outcomes, w probabilities

$$p_1, p_2, \dots, p_r$$

$$p_1 + p_2 + \dots + p_r = 1$$

Perform n trials.

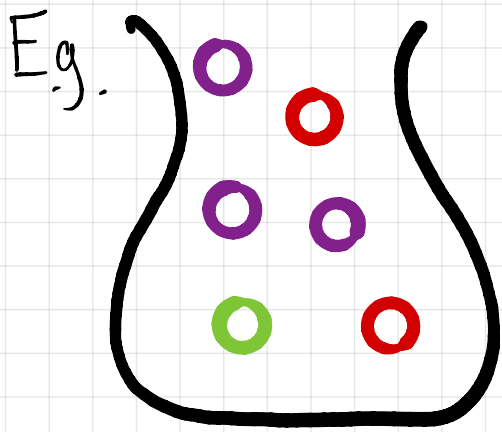
For $1 \leq j \leq r$, $X_j = \#$ times we get outcome j .

Possible values for $\underline{X} = (X_1, \dots, X_r)$: $(k_1, \dots, k_r) \leftarrow \begin{matrix} k_j \in \{0, 1, \dots, n\} \\ k_1 + \dots + k_r = n \end{matrix}$

$$P(\underline{X} = \underline{k}) = \binom{\# \text{arrangements}}{n \text{ } k_1 \text{'s, } k_2 \text{'s, } \dots, k_r \text{'s}} p_1^{k_1} p_2^{k_2} \dots p_r^{k_r} \quad (r=2): \left(\frac{n!}{k_1! k_2!} = \frac{n!}{k_1! (n-k_1)!} = \binom{n}{k_1} \right)$$

pmf of $\text{Mult}(n; p_1, \dots, p_r)$

$$\binom{n}{k_1, k_2, \dots, k_r} = \frac{n!}{k_1! k_2! \dots k_r!}$$



Sample 10 times with replacement.

$$P(3 \text{ green}, 2 \text{ red}, 5 \text{ blue})$$

$G = \# \text{green}, R = \# \text{reds}, B = \# \text{blues}$

$$\underline{X} = (G, R, B) \sim \text{Mult}(10; \frac{1}{6}, \frac{1}{3}, \frac{1}{2})$$

$$P(\underline{X} = (3, 2, 5)) = \frac{10!}{3!2!5!} \left(\frac{1}{6}\right)^3 \left(\frac{1}{3}\right)^2 \left(\frac{1}{2}\right)^5 \doteq 4.05\%$$

Note: if $r=2$, $\text{Mult}(n; p, q)$

Eg. Suppose $\underline{X} \sim \text{Mult}(n; p_1, p_2, \dots, p_r)$. Find the marginal distribution of X_1 .

$$P(X_1 = t) = \sum_{\substack{k_2, k_3, \dots, k_r \geq 0 \\ t + k_2 + \dots + k_r = n}} \frac{n!}{t! k_2! \dots k_r!} p_1^t p_2^{k_2} \dots p_r^{k_r}$$

↑
Complicated! Instead observe:

$X_1 = \# \text{ successes in } n \text{ indep. trials}$
 where success = outcome 1 ($P = p_1$)
 failure = outcomes 2- r ($P = p_2 + \dots + p_r = 1 - p$)

$$X_1 \sim \text{Bin}(n, p)$$

Expectations

Let $\underline{X} = (X_1, \dots, X_n)$ be a (discrete) random vector with joint probability mass function $p_{\underline{X}}$.

If $g: \mathbb{R}^n \rightarrow \mathbb{R}$ is a function, $Y = g(\underline{X})$ is a random variable (still discrete).

$$\begin{aligned} \mathbb{E}(Y) &= \sum_t t \cdot P(Y=t) \\ &= \sum_{\underline{k}} g(\underline{k}) P(\underline{X}=\underline{k}) \\ &= \sum_{\underline{k}} g(\underline{k}) p_{\underline{X}}(\underline{k}) \end{aligned}$$

Eg. Toss a fair coin twice, $X_1, X_2 \in \{0, 1\}$.

$$\mathbb{E}(X_1 X_2) = \sum_{k_1, k_2} \underbrace{g(k_1, k_2)}_{k_1 k_2} \underbrace{P(X_1=k_1, X_2=k_2)}_{1/4}$$

$$\begin{aligned} 2 &= 0 \cdot 0 \cdot \frac{1}{4} + 0 \cdot 1 \cdot \frac{1}{4} + 1 \cdot 0 \cdot \frac{1}{4} \\ &\quad + 1 \cdot 1 \cdot \frac{1}{4} \\ &= \frac{1}{4} = \mathbb{E}(X) \mathbb{E}(Y) \end{aligned}$$

$$\begin{aligned} \mathbb{E}(X) &= \frac{1}{2} \\ \mathbb{E}(Y) &= \frac{1}{2} \end{aligned}$$

Jointly Continuous Random Vectors

A random vector $\underline{X} = (X_1, \dots, X_n)$ has a pdf $f_{\underline{X}}: \mathbb{R}^n \rightarrow \mathbb{R}_+$ if, for "nice" subsets $B \subseteq \mathbb{R}^n$

~~\mathbb{R}^n~~ $P(\underline{X} \in B) = \int_B f_{\underline{X}}(x_1, \dots, x_n) dx_1 \dots dx_n.$

(We say X_1, X_2, \dots, X_n are jointly continuous.)

Properties:

- (1) $f_{\underline{X}} \geq 0$
- (2) $\int_{\mathbb{R}^n} f_{\underline{X}} = 1.$

Eg. Standard Multivariate Normal

$$f(x_1, \dots, x_n) = (2\pi)^{-n/2} e^{-\frac{(x_1^2 + \dots + x_n^2)}{2}}$$
$$\int_{\mathbb{R}^n} f(x_1, \dots, x_n) dx_1 \dots dx_n = (2\pi)^{-n/2} \int_{-\infty}^{\infty} e^{-x_1^2/2} dx_1 \cdot \int_{-\infty}^{\infty} e^{-x_2^2/2} dx_2 \dots \int_{-\infty}^{\infty} e^{-x_n^2/2} dx_n = (2\pi)^{-n/2} (\sqrt{2\pi}) (\sqrt{2\pi}) \dots (\sqrt{2\pi}) = 1$$

$e^{-x_1^2/2} e^{-x_2^2/2} \dots e^{-x_n^2/2}$