# 3.5–3.6 Expected values and variance

Prof. Tesler

Math 186 Winter 2019

Prof. Tesler

3.5-3.6 Expected values and variance

Math 186 / Winter 2019 1 / 24

# Expected Winnings in a Game

#### • A simple game:

- Pay \$1 to play once.
- Plip two fair coins.
- Win \$5 if HH, nothing otherwise.

• The payoff is 
$$X = \begin{cases} \$5 & \text{with probability } 1/4; \\ \$0 & \text{with probability } 3/4. \end{cases}$$

• The net winnings are

 $Y = X - 1 = \begin{cases} \$5 - \$1 = \$4 & \text{with probability } 1/4; \\ \$0 - \$1 = -\$1 & \text{with probability } 3/4. \end{cases}$ 

- Playing the game once is called a trial.
- Playing the game *n* times is an experiment with *n* trials.

# Expected Winnings in a Game

 If you play the game n times, the payoff will be \$5 about n/4 times and \$0 about 3n/4 times, totalling

 $5 \cdot n/4 + 0 \cdot 3n/4 = 5n/4$ 

• The expected payoff (long term average payoff) per game is obtained by dividing by *n*:

 $E(X) = \$5 \cdot 1/4 + \$0 \cdot 3/4 = \$1.25$ 

 For the expected winnings (long term average winnings), subtract off the bet:

$$E(Y) = E(X - 1) = \$4 \cdot 1/4 - \$1 \cdot 3/4 = \$0.25$$

- That's good for you and bad for the house.
- A fair game has expected winnings = \$0.
   A game favors the player if the expected winnings are positive.
   A game favors the house if the expected winnings are negative.

(Technical name for long term average)

• The expected value of a discrete random variable *X* is

$$E(X) = \sum_{x} x \cdot p_X(x)$$

• The expected value of a continuous random variable X is

$$E(X) = \int_{-\infty}^{\infty} x \cdot f_X(x) \, dx$$

- *E*(*X*) is often called the mean value of *X* and is denoted μ
   (or μ<sub>X</sub> if there is more than one random variable in the problem).
- $\mu$  doesn't have to be a value in the range of *X*. The previous example had range *X* = \$0 or \$5, and mean \$1.25.

# Expected Value of Binomial Distribution

- Consider a biased coin with probability p = 3/4 for heads.
- Flip it 10 times and record the number of heads, x<sub>1</sub>.
   Flip it another 10 times, get x<sub>2</sub> heads.
   Repeat to get x<sub>1</sub>, · · · , x<sub>1000</sub>.
- Estimate the average of  $x_1, \ldots, x_{1000}$ : 10(3/4) = 7.5(Later we'll show E(X) = np for the binomial distribution.)
- An estimate based on the pdf: About  $1000p_X(k)$  of the  $x_i$ 's equal k for each k = 0, ..., 10, so

average of 
$$x_i$$
's  $=\frac{\sum_{i=1}^{1000} x_i}{1000} \approx \frac{\sum_{k=0}^{10} k \cdot 1000 p_X(k)}{1000} = \sum_{k=0}^{10} k \cdot p_X(k)$ 

which is the formula for E(X) in this case.

## Interpretation of the word "Expected"

- Although E(X) = 7.5, this is not a possible value for X.
- Expected value does *not* mean we anticipate observing that value.
- It means the long term average of many independent measurements of X will be approximately E(X).

## Function of a Random Variable

• Let *X* be the value of a roll of a biased die and  $Z = (X - 3)^2$ .

X	$p_X(x)$	$z = (x - 3)^2$	$p_Z(z)$
1	$q_1$	4	
2	$q_2$	1	
3	$q_3$	0	$p_Z(0) = q_3$
4	$q_4$	1	$p_{Z}(1) = q_{2} + q_{4}$
5	$q_5$	4	$p_Z(4) = q_1 + q_5$
6	$q_6$	9	$p_Z(9) = q_6$

- pdf of *X*: Each  $q_i \ge 0$  and  $q_1 + \cdots + q_6 = 1$ .
- **pdf of** *Z*: Each probability is also  $\ge 0$ , and the total sum is also 1.

# Expected Value of a Function

• Let *X* be the value of a roll of a biased die and  $Z = (X - 3)^2$ .

X	$p_X(x)$	$z = (x - 3)^2$	$p_Z(z)$
1	$q_1$	4	
2	$q_2$	1	
3	$q_3$	0	$p_Z(0) = q_3$
4	$q_4$	1	$p_{Z}(1) = q_{2} + q_{4}$
5	$q_5$	4	$p_Z(4) = q_1 + q_5$
6	$q_6$	9	$p_{Z}(9) = q_{6}$

• E(Z), in terms of values of Z and the pdf of Z, is

$$E(Z) = \sum_{z} z \cdot p_{Z}(z) = 0(q_{3}) + 1(q_{2} + q_{4}) + 4(q_{1} + q_{5}) + 9(q_{6})$$

• Regroup it in terms of *X*:

$$= 4q_1 + 1q_2 + 0q_3 + 1q_4 + 4q_5 + 9q_6 = \sum_{x=1}^{6} (x-3)^2 p_X(x)$$

# Expected Value of a Function

- Let *X* be a discrete random variable, and g(X) be a function, such as  $(X-3)^2$ .
- The expected value of g(X) is

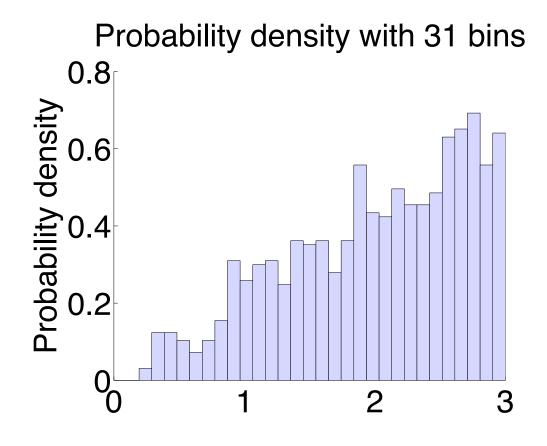
$$E(g(X)) = \sum_{x} g(x) p_{X}(x)$$

For a continuous random variable,

$$E(g(X)) = \int_{-\infty}^{\infty} g(x) f_X(x) \, dx$$

• Note that if Z = g(X) then E(Z) = E(g(X)).

# Expected Value of a Continuous distribution



 Consider the dartboard of radius 3 example, with pdf

 $f_{R}(r) = \begin{cases} 2r/9 & \text{if } 0 \leqslant r \leqslant 3; \\ 0 & \text{otherwise.} \end{cases}$ 

- Throw n darts and make a histogram with k bins.
- $r_1, r_2, \ldots$  are representative values of *R* in each bin.
- The bin width is  $\Delta r = 3/k$ , the height is  $\approx f_R(r_i)$ , and the area is  $\approx f_R(r_i) \Delta r$ .
- The approximate number of darts in bin *i* is  $nf_R(r_i) \Delta r$ .

• The estimated average radius is

$$\frac{\sum_{i} r_{i} \cdot n f_{R}(r_{i}) \Delta r}{n} = \sum_{i} r_{i} \cdot f_{R}(r_{i}) \Delta r$$

• As  $n, k \to \infty$ , the histogram smoothes out and this becomes

$$\int_0^3 r \cdot f_R(r) \, dr$$

## Mean of a continuous distribution

• Consider the dartboard of radius 3 example, with pdf

$$f_{R}(r) = \begin{cases} 2r/9 & \text{if } 0 \leqslant r \leqslant 3; \\ 0 & \text{otherwise.} \end{cases}$$

The "average radius" (technically the mean radius or expected value of *R*) is

$$\mu = E(R) = \int_{-\infty}^{\infty} r \cdot f_R(r) \, dr = \int_0^3 r \cdot \frac{2r}{9} \, dr = \int_0^3 \frac{2r^2}{9} \, dr$$
$$= \left. \frac{2r^3}{27} \right|_0^3 = \frac{2(3^3 - 0^3)}{27} = 2$$

## **Expected Values** — Properties

• The gambling slide earlier had E(X - 1) = E(X) - 1.

#### Theorem

E(aX+b) = aE(X) + b where a, b are constants.

#### Proof (discrete case).

E

$$(aX+b) = \sum_{x} (ax+b) \cdot p_{X}(x)$$
$$= a \sum_{x} x \cdot p_{X}(x) + b \sum_{x} p_{X}(x)$$
$$= a \cdot E(X) + b \cdot 1 = aE(X) + b$$

## **Expected Values** — Properties

• The gambling slide earlier had E(X - 1) = E(X) - 1.

#### Theorem

E(aX+b) = aE(X) + b where a, b are constants.

#### Proof (continuous case).

$$E(aX+b) = \int_{-\infty}^{\infty} (ax+b) \cdot f_X(x) dx$$
  
=  $a \int_{-\infty}^{\infty} x \cdot f_X(x) dx + b \int_{-\infty}^{\infty} f_X(x) dx$   
=  $a \cdot E(X) + b \cdot 1 = aE(X) + b$ 

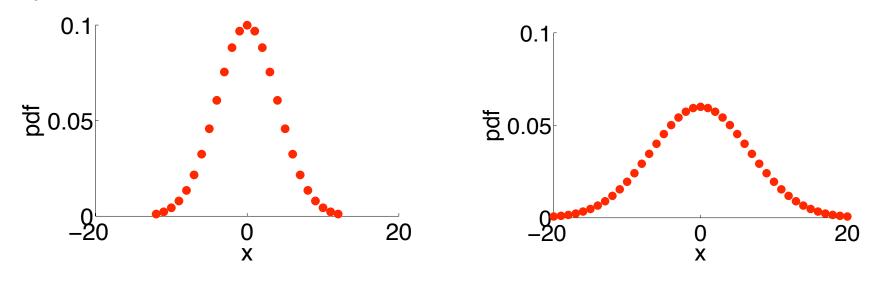
# These properties hold for both discrete and continuous random variables:

- E(aX+b) = a E(X) + b for any constants a, b.
- E(aX) = a E(X)
- E(b) = b

• 
$$E(g(X) + h(X)) = E(g(X)) + E(h(X))$$

## Variance

 These distributions both have mean=0, but the right one is more spread out.



• The variance of *X* measures the square of the spread from the mean:

$$\sigma^2 = \operatorname{Var}(X) = E((X - \mu)^2)$$

• The standard deviation of *X* is  $\sigma = \sqrt{Var(X)}$  and measures how wide the curve is.

- The same definition  $Var(X) = E((X \mu)^2)$  is used for both the discrete and continuous cases, but expected value is computed differently in the two cases.
- Why don't we use  $E(X \mu)$  or  $E(|X \mu|)$  to measure the spread?
  - $E(X \mu) = E(X) \mu = \mu \mu = 0$ , so it doesn't measure the spread.
  - Both |X μ| and (X μ)<sup>2</sup> are nonnegative.
     We will see that E((X μ)<sup>2</sup>) leads to useful properties.
     It turns out that E(|X μ|) does not have nice properties.

# Variance formula $\sigma^2 = E((X - \mu)^2)$

- Consider the dartboard of radius 3 example, with pdf  $f_{R}(r) = \begin{cases} 2r/9 & \text{if } 0 \leqslant r \leqslant 3; \\ 0 & \text{otherwise.} \end{cases}$
- $\mu = 2$  from earlier slide.

• 
$$\sigma^2 = \operatorname{Var}(R) = E((R - \mu)^2) = E((R - 2)^2)$$
  

$$= \int_{-\infty}^{\infty} (r - 2)^2 f_R(r) \, dr = \int_0^3 \frac{(r - 2)^2 \cdot 2r}{9} \, dr$$

$$= \int_0^3 \frac{2r^3 - 8r^2 + 8r}{9} \, dr = \left(\frac{r^4}{18} - \frac{8r^3}{27} + \frac{4r^2}{9}\right) \Big|_0^3$$

$$= \left(\frac{3^4}{18} - \frac{8(3^3)}{27} + \frac{4(3^2)}{9}\right) - 0 = \frac{1}{2}$$

• Variance:  $\sigma^2 = \frac{1}{2}$  Standard deviation:  $\sigma = \sqrt{1/2}$ 

## Variance — Second formula

#### There are two equivalent formulas to compute variance. In any problem, choose the easier one:

$$\sigma^2 = \operatorname{Var}(X) = E((X - \mu)^2)$$
 (Definition)  
=  $E(X^2) - \mu^2$  (Sometimes easier to compute)

#### Proof.

$$Var(X) = E((X - \mu)^{2})$$
  
=  $E(X^{2} - 2\mu X + \mu^{2})$   
=  $E(X^{2}) - 2\mu E(X) + \mu^{2}$   
=  $E(X^{2}) - 2\mu^{2} + \mu^{2}$   
=  $E(X^{2}) - \mu^{2}$ 

Consider the dartboard of radius 3 example, with pdf

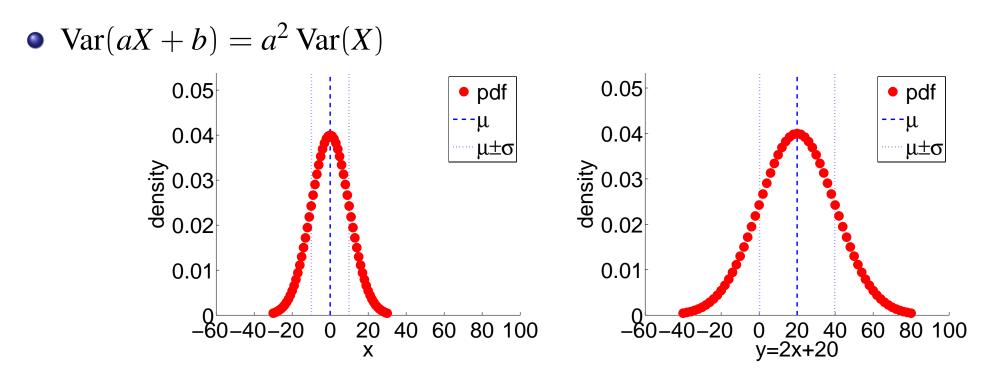
$$f_{R}(r) = \begin{cases} 2r/9 & \text{if } 0 \leqslant r \leqslant 3; \\ 0 & \text{otherwise.} \end{cases}$$

• 
$$\mu = E(R) = 2$$

• 
$$E(R^2) = \int_0^3 r^2 \cdot \frac{2r}{9} dr = \int_0^3 \frac{2r^3}{9} dr = \frac{2r^4}{36} \Big|_0^3 = \frac{2(81-0)}{36} = \frac{9/2}{36}$$

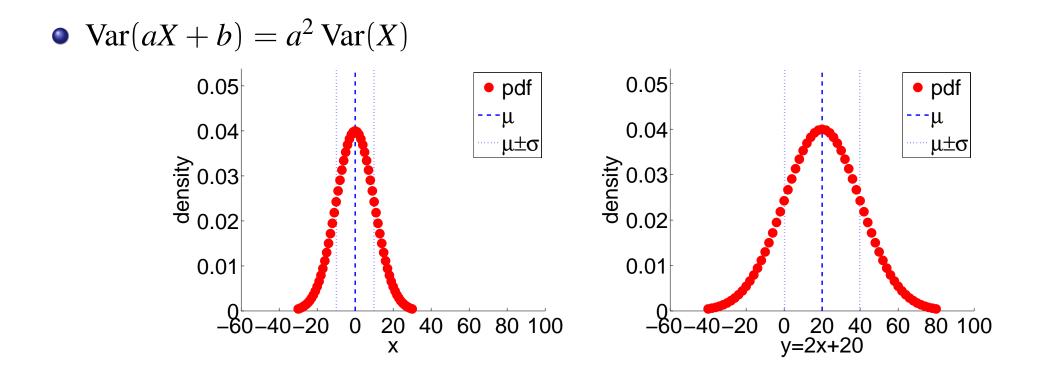
• Variance:  $\sigma^2 = E(R^2) - \mu^2 = \frac{9}{2} - 2^2 = \frac{1}{2}$ Standard deviation:  $\sigma = \sqrt{1/2}$ 

# Variance — Properties



- Adding b shifts the curve without changing the width, so b disappears on the right side of the formula.
- Multiplying by *a* dilates the width a factor of *a*, so variance goes up a factor  $a^2$ .
- For Y = aX + b, we have  $\sigma_Y = |a| \sigma_X$ .
- **Example:** Convert measurements in  $^{\circ}C$  to  $^{\circ}F$ : F = (9/5)C + 32  $\mu_F = (9/5)\mu_C + 32$   $\sigma_F = (9/5)\sigma_C$

## Variance — Properties



#### Proof of $Var(aX + b) = a^2 Var(X)$ .

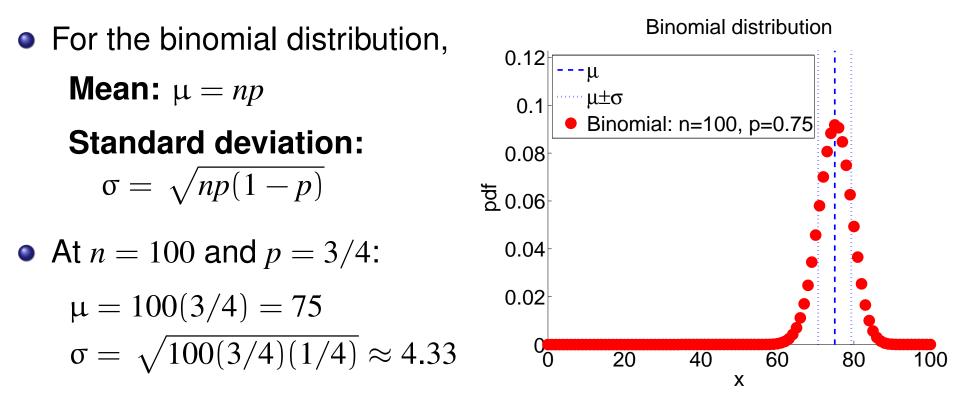
$$E((aX + b)^{2}) = E(a^{2}X^{2} + 2abX + b^{2}) = a^{2}E(X^{2}) + 2abE(X) + b^{2}$$
  

$$(E(aX + b))^{2} = (aE(X) + b)^{2} = a^{2}(E(X))^{2} + 2abE(X) + b^{2}$$
  

$$Var(aX + b) = difference = a^{2}\left(E(X^{2}) - (E(X))^{2}\right)$$
  

$$= a^{2}Var(X) \square$$

# Mean and Variance of the Binomial Distribution



Approximately 68% of the probability is for X between μ±σ.
 Approximately 95% of the probability is for X between μ±2σ.
 More on that in Chapter 4.

# Mean of the Binomial Distribution

### Proof that $\mu = np$ for binomial distribution.

$$E(X) = \sum_{k} k \cdot p_X(k)$$
  
=  $\sum_{k=0}^{n} k \cdot {n \choose k} p^k q^{n-k}$ 

$$\begin{array}{ll} \textbf{Calculus Trick:} & (p+q)^n = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k} \\ \text{Differentiate:} & \frac{\partial}{\partial p} (p+q)^n = \sum_{k=0}^n k\binom{n}{k} p^{k-1} q^{n-k} \\ \text{Times } p: & p \frac{\partial}{\partial p} (p+q)^n = \sum_{k=0}^n k\binom{n}{k} p^k q^{n-k} = E(X) \\ \text{Evaluate left side:} & p \frac{\partial}{\partial p} (p+q)^n = p \cdot n(p+q)^{n-1} \\ & = p \cdot n \cdot 1^{n-1} = np \quad \text{since } p+q = 1. \end{array}$$

So E(X) = np.

We'll do  $\sigma = \sqrt{np(1-p)}$  later.