

# NOTES ON RANDOM WALKS

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## 1. SIMPLE RANDOM WALKS

Let  $\xi$  be a random variable taking integer values, and  $x \in \mathbb{Z}$ . For  $i = 1, 2, \dots$  we successively walk  $\xi$  units right or left, depending on the sign of  $\xi$ , from an initial position  $x$ . We define a sequence of random variables  $X_n$ , expressing the position at time  $n$  by

$$(1) \quad X_0 = x, X_1 = x + \xi_1, \dots, X_n = x + \xi_1 + \xi_2 + \dots + \xi_n,$$

where the  $\xi_i$  are independent, identically distributed with distribution the same as that of  $\xi$ , with finite mean  $\mu$  and finite variance  $\sigma^2$ . The sequence  $\{X_n\}$  is called a *random walk*. If  $\xi$  takes only the values  $0, -1, +1$ , the sequence is a *simple random walk*. By the additivity of expectation of independent random variables,

$$(2) \quad E[X_n] = x + nE[\xi].$$

Now let  $T$  be a random variable taking positive integer values, with finite mean  $E[T]$ , independent of the  $\xi_i$  for  $i > T$ . We think of  $T$  as a stopping time, and are interested in the random variable  $X_T$  (which is a *random sum*).

**1.1. Wald's identity.** . The expected value  $E[X_T]$  and variance  $\text{Var}[X_T]$  satisfy the following identities due to Wald.

$$(3) \quad E[X_T] = x + E[\xi]E[T]$$

If, in addition,  $E[\xi] = 0$ , then

$$(4) \quad \text{Var}[X_T] = \sigma^2 E[T],$$

where  $\sigma^2 = \text{Var}(\xi_i)$ . Idea of proof:

(5)

$$E[X_T] = \sum_{n=1}^{\infty} E[X_T|T = n]P\{T = n\}$$

(6) 
$$= \sum_{n=1}^{\infty} E[X_n]P\{T = n\} \quad (\text{by independence of } \xi_i \text{ and } T \text{ for } i > T)$$

(7) 
$$= \sum_{n=1}^{\infty} (x + nE[\xi])P\{T = n\}$$

(8) 
$$= x + E[\xi] \sum_{n=1}^{\infty} nP\{T = n\},$$

which proves (3). For (4), assume  $E[\xi] = 0$  and write

(9) 
$$\text{Var}(X_T) = E[(\sum_{i=1}^T \xi_i - E[\xi]E[T])^2] = E[(\sum_{i=1}^T \xi_i)^2]$$

By the law of total probability,

(10)

$$E[(\sum_{i=1}^T \xi_i)^2] = \sum_{n=1}^{\infty} E[(\sum_{i=1}^T \xi_i)^2|T = n] \cdot P\{T = n\} = \sum_{n=1}^{\infty} E[(\sum_{i=1}^n \xi_i)^2] \cdot P\{T = n\}$$

By independence of the  $\xi_j$ , and the assumption  $E[\xi_j] = 0$ ,

(11) 
$$E[\xi_j \xi_k] = 0 \quad \text{for } j \neq k.$$

Then

(12) 
$$\begin{aligned} \text{Var}(X_T) &= \sum_{n=1}^{\infty} (\sum_{j=1}^n E[\xi_j^2]) \cdot P\{T = n\} = \\ &= \text{Var}[\xi] (\sum_{n=1}^{\infty} nP\{T = n\}) = \text{Var}[\xi]E[T] \end{aligned}$$

## 1.2. Applications of Wald's identity to simple random walks

**(gambler's ruin).** In the following examples, players Jack and Jill will start with \$5 and \$10 respectively and play a game by making a series of \$1 bets until one of them loses all his/her money. We'll assume that in each bet, Jack wins with probability  $p$ , Jill with probability  $q$ , and tie (no money exchanged) with probability  $r$ , so that

$$p + q + r = 1.$$

To model this as a simple random walk, we let  $\xi_1, \xi_2, \dots$  be a sequence of i.i.d. random variables taking the values  $1, -1, 0$  with probabilities  $p, q, r$  respectively. Let  $X_n$  be the amount of money that Jack has after  $n$  bets, so that

$$X_n = 5 + \xi_1 + \xi_2 + \dots + \xi_n.$$

Note that by the rules of the game,  $0 \leq X_n \leq 15$  for all  $n$ . Let  $T$  be the number of bets made until the game ends. Then  $T$  is a stopping time, and  $X_T$  represents the amount of money that Jack has when the game ends. The random sum  $X_T$  can take only two values, 0 (Jill wins) or 15 (Jack wins). We will use Wald's identities to calculate  $E[T]$  and the respective probabilities of Jack or Jill winning.

**Example 1.1.** Suppose that  $p = q = 1/2$ ,  $r = 0$ , so that  $E[\xi] = 0$ . By the first Wald identity (3), it follows that

$$E[X_T] = 5.$$

Then

$$(13) \quad E[X_T] = 0 \cdot P(X_T = 0) + 15 \cdot P(X_T = 15) = 5,$$

so that

$$P(X_T = 15) = 1/3,$$

i.e.

$$P(\text{Jack wins}) = 1/3,$$

while

$$P(\text{Jill wins}) = 2/3.$$

This also gives

$$(14) \quad \begin{aligned} \text{Var}[X_T] &= E[X_T^2] - 5^2 = 0 \cdot P(X_T^2 = 0) + 15^2 \cdot P(X_T^2 = 15^2) - 25 \\ &= 15^2 \cdot 1/3 - 25 = 50. \end{aligned}$$

By the second Wald identity (4),  $\text{Var}[X_T] = E[\xi^2]E[T]$ . Also,

$$\xi^2 \equiv 1 \implies E[\xi^2] = 1.$$

Hence  $E[T] = 50$ .

**Example 1.2.** Now suppose  $p \neq q$ , so that (4) does not hold. We set up an inductive process for  $0 \leq x \leq 15$  by putting

$$(15) \quad X_n(x) = x + \sum_{j=1}^n \xi_j, \quad X_0 = x, \quad \text{and let } f(x) := P(X_T(x) = 15)$$

Then

$$(16) \quad f(0) = 0, \quad f(15) = 1,$$

and we are interested in computing  $f(5)$ . Conditioning on the possible values of  $\xi_1$ , we have

$$(17) \quad f(x) = f(x+1)\mathbb{P}(\xi_1 = 1) + f(x-1)\mathbb{P}(\xi_1 = -1) \\ = p \cdot f(x+1) + q \cdot f(x-1).$$

Writing

$$f(x) = pf(x) + qf(x),$$

we obtain from (17),

$$(18) \quad f(x) - f(x-1) = (p/q)(f(x+1) - f(x))$$

Now define

$$h(x) := f(x) - f(x-1),$$

and note that  $h(1) = f(1)$ . The collapsing sum gives

$$(19) \quad f(x) = \sum_{j=1}^x h(j) + f(0) = \sum_{j=1}^x h(j).$$

On the other hand, since

$$h(x+1) = (q/p)h(x)$$

by (18), we have by induction,

$$h(x) = (q/p)^{x-1}h(1) = (q/p)^{x-1}f(1).$$

From (16) and (19), we then have

$$1 = f(15) = \sum_{j=1}^{15} h(j) = \sum_{j=1}^{15} (q/p)^{j-1}f(1),$$

which gives

$$f(1) = \frac{1}{\sum_{j=1}^{15} (q/p)^{j-1}}$$

By putting this into (19), we may calculate the desired  $f(5)$  explicitly by

$$(20) \quad f(5) = \sum_{j=1}^5 h(j)f(1) = \sum_{j=1}^5 (q/p)^{j-1}f(1) = \\ \frac{\sum_{j=1}^5 (q/p)^{j-1}}{\sum_{j=1}^{15} (q/p)^{j-1}} = \frac{1 - (q/p)^5}{1 - (q/p)^{15}}$$

To find  $E[T]$ , we begin by finding  $E[X_T]$  in order to apply (3).

$$(21) \quad E[X_T] = 0 \cdot [\mathbb{P}\{X_T = 0\}] + 15f(5) = 15 \frac{1 - (q/p)^5}{1 - (q/p)^{15}}$$

Then

$$(22) \quad E[T] = (E[X_T] - 5)/E[\xi] = \frac{15 \frac{1-(q/p)^5}{1-(q/p)^{15}} - 5}{p - q}.$$