

# MATH180C: Introduction to Stochastic Processes II

Lecture A00: [math-old.ucsd.edu/~ynemish/teaching/180cA](http://math-old.ucsd.edu/~ynemish/teaching/180cA)

Lecture B00: [math-old.ucsd.edu/~ynemish/teaching/180cB](http://math-old.ucsd.edu/~ynemish/teaching/180cB)

Today: FSA for general MC

Next: PK 6.3, 6.6, Durrett 4.2

Week 3:

- homework 2 (due Friday April 15)
- HW 1 regrades: Wednesday April 13

## Q-matrices and Markov chains (cont.)

$P(t)$  satisfies properties (a)-(d) from Theorem A.

$\Rightarrow$  there is a Q-matrix  $Q$  such that

$$P(t) = e^{tQ}$$

$$P_{ij}(h) = q_{ij}h + o(h) \quad i \neq j$$

In particular,

$$P_{ii}(h) = 1 + q_{ii}h + o(h)$$

$$P(h) = I + Qh + o(h) \quad \text{as } h \rightarrow 0$$

This implies the one-to-one correspondance between Q-matrices and continuous time MC with right-continuous sample paths.

$Q$  is called the infinitesimal generator of  $(X_t)_{t \geq 0}$

## Sojourn time description

Let  $Q = (q_{ij})_{i,j=0}^N$  be a Q-matrix. Denote  $q_i = \sum_{j \neq i} q_{ij}$

so that

$$Q = \begin{pmatrix} -q_0 & q_{01} & q_{02} & \cdots \\ q_{10} & -q_1 & q_{12} & \cdots \\ q_{20} & q_{21} & -q_2 & \cdots \\ \vdots & \vdots & \ddots & \vdots \end{pmatrix} \quad q_0 = \sum_{i \neq 0} q_{0i}$$

Denote  $Y_k := X_{W_k}$  (jump chain).

Then the MC with generator matrix  $Q$  has the following equivalent jump and hold description

- sojourn times  $S_k$  are independent r.v.

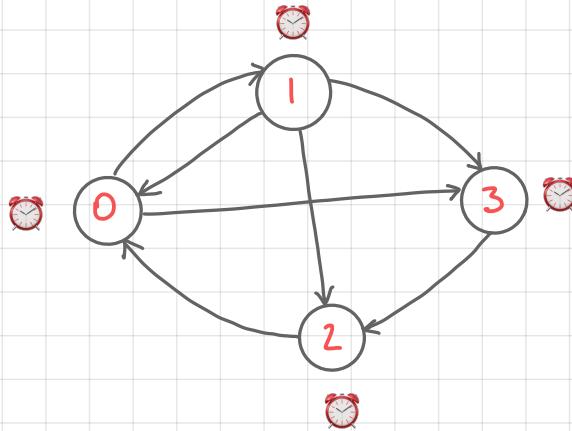
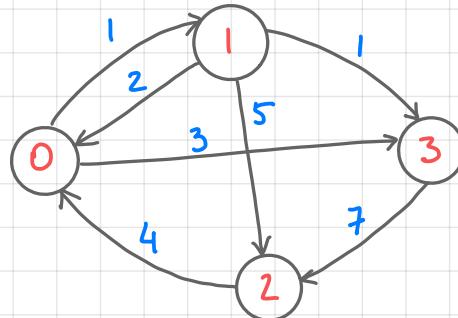
with  $P(S_k > t \mid Y_k = i) = e^{-q_i t} \quad (S_k \sim \text{Exp}(q_i))$

- transition probabilities  $P(Y_{k+1} = j \mid Y_k = i) = \frac{q_{ij}}{q_i}$

## Example

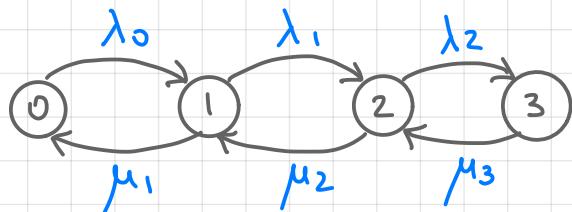
	0	1	2	3
0	-4	1	0	3
1	2	-7	5	1
2	4	0	-4	0
3	0	0	7	-7

$i \xrightarrow{\alpha} j =$



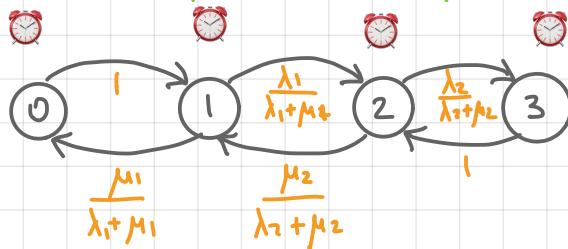
## Example

Birth and death process on  $\{0, 1, 2, 3\}$



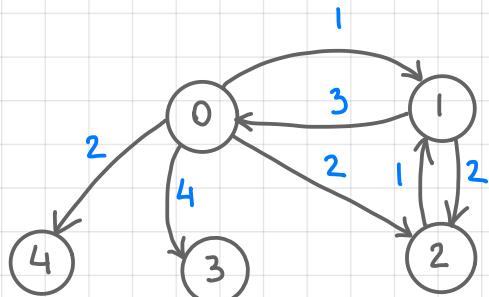
$$Q = \begin{pmatrix} -\lambda_0 & \lambda_0 & 0 & 0 \\ \mu_1 & -(\lambda_1 + \mu_1) & \lambda_1 & 0 \\ 0 & \mu_2 & -(\lambda_2 + \mu_2) & \lambda_2 \\ 0 & 0 & \mu_3 & -\mu_3 \end{pmatrix}$$

$\text{Exp}(\lambda_0) \quad \text{Exp}(\lambda_1 + \mu_1) \quad \text{Exp}(\lambda_2 + \mu_2) \quad \text{Exp}(\mu_3)$



## General continuous time finite state MCs

### Rate diagram



### Generator

$$Q = \begin{pmatrix} 0 & 1 & 2 & 3 & 4 \\ -9 & 1 & 2 & 4 & 2 \\ 3 & -5 & 2 & 1 & -1 \\ 2 & 1 & -1 & 0 & 0 \\ 3 & 0 & 0 & 0 & 0 \end{pmatrix}$$

### Infinitesimal description

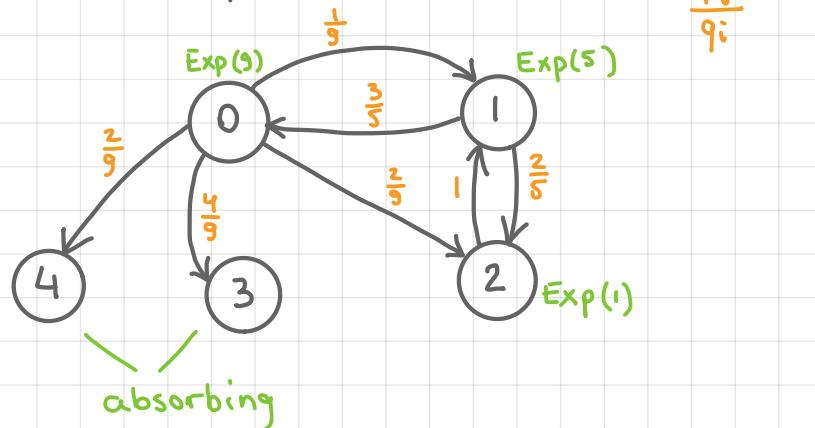
$$P_{ij}(h) = q_{ij}h + o(h), i \neq j$$

$$P_{ii}(h) = 1 - q_{ii}h + o(h)$$

$$P_{02}(h) = 2h + o(h)$$

$$P_{00}(h) = 1 - 9h + o(h)$$

### Jump and hold



## Absorption probabilities for finite state chains

By considering the jump chain  $(Y_n)_{n \geq 0}$  with  $Y_n = X_{w_n}$  and its transition probabilities  $P(Y_{n+1}=j | Y_n=i) = \frac{q_{ij}}{q_i}$  we can apply the first step analysis to compute, e.g., the absorption probabilities (similarly as for B&D)

If state  $i$  is absorbing, then  $q_{ij} = 0$  for all  $j \neq i$  (no jumps from state  $i$ ), so  $q_i = q_{ii} = 0$ . Let  $Q$  be given by

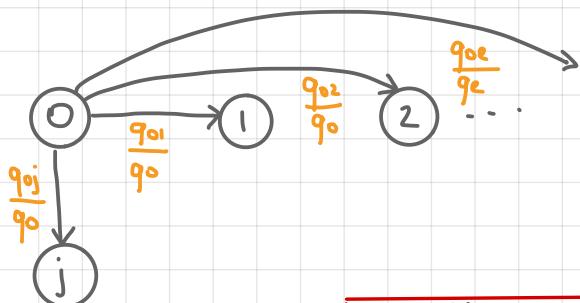
$$Q = \begin{array}{c|cc|c} & 0 & \cdots & K-1 & K \cdots N \\ \hline 0 & -q_0 & & \cdots & q_{ij} \\ \vdots & \vdots & & & \vdots \\ K-1 & q_{ij} & \cdots & -q_{K-1} & 0 \\ K & & & & 0 \\ \vdots & & & & \ddots \\ N & 0 & & & 0 \end{array}$$

with  $\{0, \dots, K-1\}$  transient,  
 $\{K, \dots, N\}$  absorbing

## Absorption probabilities for finite state chains

$$Q = \begin{pmatrix} 0 & \cdots & k-1 & k \cdots N \\ -q_0 & & & \\ \vdots & & \ddots & q_{ij} \\ q_{ij} & \cdots & -q_{k-1} & \\ \vdots & & 0 & \\ K & & & 0 \\ \vdots & & & \ddots \\ N & & & 0 \end{pmatrix}$$

Jump chain



Let  $i \in \{0, \dots, k-1\}$ ,  $j \in \{k, \dots, N\}$ .

Let  $M = \min\{n : Y_n \in \{k, \dots, N\}\}$

Denote  $u_i^{(j)} = P(Y_M = j | X_0 = i)$ .

Then FSA leads to the system

$$u_i^{(j)} = P(Y_M = j | Y_0 = i)$$

=

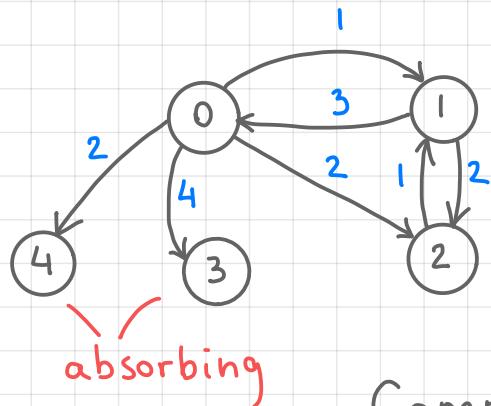
$$u_i^{(j)} = \frac{q_{ij}}{q_i} + \sum_{\substack{c=0 \\ c \neq i}}^{k-1} \frac{q_{ic}}{q_i} u_c^{(j)}$$

$P(Y_{n+1} = j | Y_n = i)$

$P(Y_{n+1} = c | Y_n = i)$

## Example

### Rate diagram



### Generator

$$Q = \begin{pmatrix} 0 & 1 & 2 & 3 & 4 \\ 1 & -9 & 1 & 2 & 4 \\ 2 & 3 & -5 & 2 & 1 \\ 3 & 1 & 1 & -1 & 0 \\ 4 & 0 & 0 & 0 & 0 \end{pmatrix}$$

Compute  $P(Y_{N=3})$  if  $P(X_0=i) = p_i$  for  $i=0,1,2$

Denote  $u_i = P(Y_{N=3} | Y_0=i)$ .

$$\sum p_i = 1$$

$$\left\{ \begin{array}{l} u_0 = \\ u_1 = \\ u_2 = \end{array} \right.$$

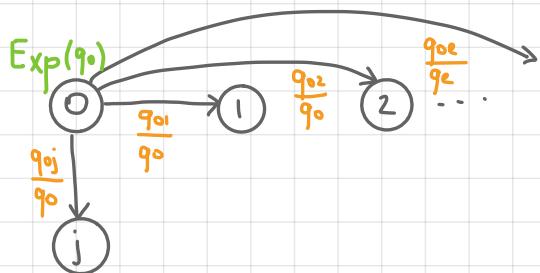
$$\left\{ \begin{array}{l} \\ \\ u_2 = u_1 \end{array} \right.$$

$$P(Y_{N=3}) =$$

## Mean time to absorption

Similar analysis as was applied to B&D processes can be used to compute the mean time to absorption: before each jump from step  $i$  to state  $j$  the process sojourns on average in state  $i$ .

$$Q = \begin{matrix} & \begin{matrix} 0 & \cdots & K-1 & K & \cdots & N \end{matrix} \\ \begin{matrix} 0 \\ \vdots \\ K-1 \\ K \\ \vdots \\ N \end{matrix} & \left( \begin{array}{cccccc} -q_0 & & & & q_{ij} & \\ \vdots & & & & \vdots & \\ q_{ij} & \cdots & -q_{k-1} & & & \\ & 0 & & & 0 & \\ & & & \ddots & & \\ & & & & 0 & \end{array} \right) \end{matrix}$$



$$\text{Let } T = \min\{t : X_t \in \{K, \dots, N\}\}$$

$$M = \min\{n : Y_n \in \{K, \dots, N\}\}$$

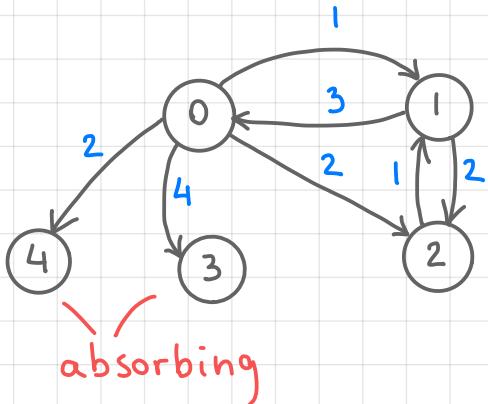
Denote  $w_i =$

Then FSA gives

$$w_i =$$

## Example

### Rate diagram



### Generator

$$Q = \begin{pmatrix} 0 & 1 & 2 & 3 & 4 \\ -9 & 1 & 2 & 4 & 2 \\ 3 & -5 & 2 & 0 & 0 \\ 1 & 1 & -1 & 0 & 0 \\ 2 & & & 0 & 0 \end{pmatrix}$$

$$T = \min \{ t : X_t \in \{3, 4\} \}$$

$$w_i = E(T | X_0 = i)$$

$$\left\{ \begin{array}{l} w_0 = \\ w_1 = \\ w_2 = \end{array} \right.$$

$$\left\{ \begin{array}{l} \\ \\ w_2 = 1 + w_1 \end{array} \right.$$

## Kolmogorov equations

Jump and hold description is very intuitive, gives a very clear picture of the process, but does not answer to some very basic questions, e.g., computing  $P_{ij}(t) := P(X_t = j | X_0 = i)$ .

For computing the transition probabilities the differential equation approach is more appropriate.

In order to derive the system of differential equations for  $P_{ij}(t)$  from the infinitesimal description, we start from the familiar relation:

Chapman - Kolmogorov equation (semigroup property)

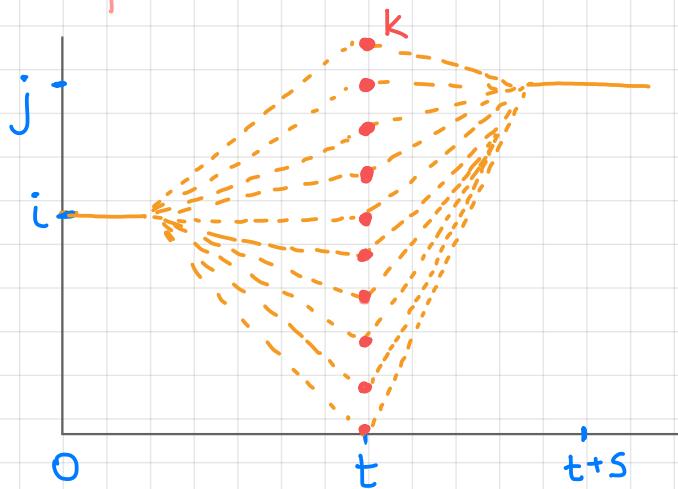
## Chapman-Kolmogorov equation

$$P_{ij}(t+s) = P(X_{t+s} = j \mid X_0 = i) \quad \text{condition on the value of } X_t$$

=

Markov =

stationary  
trans. prob. =



Or in matrix form

## Kolmogorov forward equations

Apply Chapman-Kolmogorov equations to compute

$P_{ij}(t+h) :$

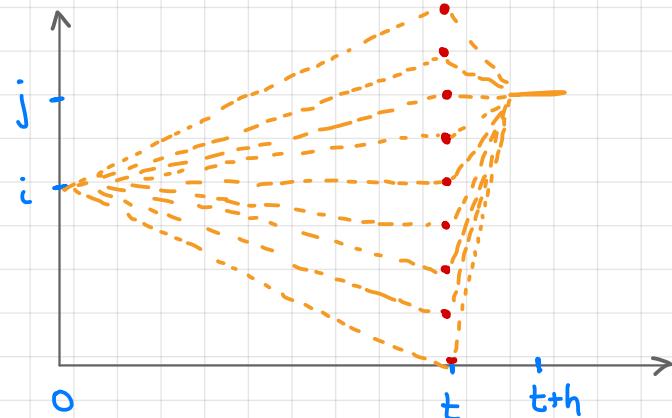
$P_{ij}(t+h) =$

Use infinitesimal description:

$$P_{kj}(h) = \begin{cases} q_{kj} h + o(h), & k \neq j \\ 1 + q_{jj} h + o(h), & k = j \end{cases}$$

(\*) =

=



$$\frac{d}{dt} P(t) = P(t) Q$$

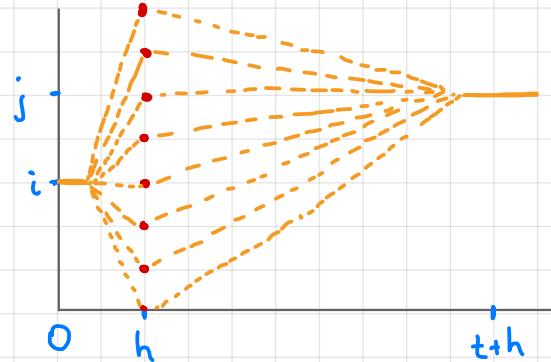
## Kolmogorov backward equations

$$P_{ij}(t+h) = \sum_{k=0}^N P_{ik}(h) P_{kj}(t)$$

$$= (1 + q_{ii} h + o(h)) P_{ij}(t)$$

$$+ \sum_{\substack{k=0 \\ k \neq i}}^N (q_{ik} h + o(h)) P_{kj}$$

$$= P_{ij}(t) + \sum_{k=0}^N q_{ik} P_{kj}(t) h + o(h)$$



## Kolmogorov equations. Remarks

1.  $e^{tQ}$  satisfies both (forward and backward) equations.  
Indeed, omitting technical details, differentiate term-by-term

$$\frac{d}{dt} e^{tQ} = \frac{d}{dt} \left( \sum_{k=0}^{\infty} \frac{Q^k t^k}{k!} \right) =$$

$$\text{Now } \sum_{k=1}^{\infty} \frac{Q^k}{(k-1)!} t^{k-1} = \sum_{\ell=0}^{\infty} \frac{Q^{\ell+1}}{\ell!} t^{\ell} =$$

2. Redundancy is related to the stationarity of transition probabilities. If transition probabilities

$$P_{ij}(s,t) = P(X_t=j | X_s=i)$$
 are not stationary, then

$\frac{\partial}{\partial t} P_{ij}(s, t)$   $\rightarrow$  forward equation ,  $\frac{\partial}{\partial s} P_{ij}(s, t)$   $\rightarrow$  backward equation