

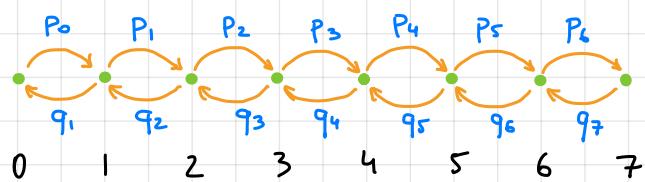
MATH 285: Stochastic Processes

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Today: Positive and null recurrence

- Homework 2 is due on Friday, January 21 11:59 PM

Birth and death processes (infinite state space)



$$S = \{0, 1, 2, 3, \dots\}$$

$$p(i, i+1) = p_i, \quad p(i, i-1) = 1 - p_i$$

$$p(0, 1) = p_0, \quad p(0, 0) = 1 - p_0$$

$p_0 \in [0, 1]$, $p_0 = 0$ absorbing, $p_0 = 1$ reflecting

Model of population growth : X_n = size of the population at time n

$P_i[\exists_{n \geq 0} : X_n = 0]$ – extinction probability

$P_i[X_n \rightarrow \infty \text{ as } n \rightarrow \infty]$ – probability that population explodes

Denote $h(i) := P_i[\exists_{n \geq 0} X_n = 0] = P_i[\tau_0 < \infty]$, $\tau_0 = \min\{n \geq 0 : X_n = 0\}$

First step analysis:

Theorem 7.0 $(h(0), h(1), \dots)$ is the minimal solution to

$$\begin{cases} h(0) = 1 \\ h(i) = \sum_{j=0}^{\infty} p(i, j) h(j) \end{cases}$$

Positive and null recurrence

Let (X_n) be a Markov chain, and let i be a recurrent state. Starting from i , (X_n) revisits i infinitely many times, $\mathbb{P}_i[X_n=i \text{ for infinitely many } n] = 1$

How often does (X_n) revisit state i ?

(i) After n steps, (X_n) revisits $i \approx \frac{n}{2}$ times, spends half of the time at i

(ii) After n steps, (X_n) revisits $i \approx \sqrt{n}$ times, the fraction of time spent at i tends to 0 as $n \rightarrow \infty$, $\frac{\sqrt{n}}{n} \rightarrow 0, n \rightarrow \infty$

Def 9.2 Let i be a recurrent state for MC (X_n) .

Denote $T_i = \min\{n \geq 1 : X_n = i\}$. If , then we call i
. If , then we call i .

Positive and null recurrence

Remark If i is recurrent, then $P_i[T_i] < \infty$. But it is still possible that $E[T_i] = \infty$ or that $E[T_i] < \infty$.

Example: $Y_1, Y_2 \in \mathbb{N}$, $P[Y_1 = k] = \dots$, $Y_2 = \dots$, $P[Y_2 = 2^k] = \dots$

$$P[Y_1 < \infty] = P[Y_2 < \infty] = 1, E[Y_1] = \dots, E[Y_2] = \dots$$

Prop 9.4 In a finite-state irreducible Markov chain all states are

Proof. Fix state $j \in S$

(1) There exist $N \in \mathbb{N}$ and $q \in (0, 1)$ such that for any $i \in S$ (probability of reaching j from i in the next N steps)

Since (X_n) is irreducible,

Take

Positive and null recurrence

(2) For any $i \in S$ $\mathbb{P}_i[T_j > N] \leq$. | follows from (1)

(3) For any $k \in \mathbb{N}$, $\mathbb{P}_j[T_j > (k+1)N] \leq$

For any $i \in S$ $\mathbb{P}_j[T_j > (k+1)N \mid T_j > kN, X_{kN} = i] =$
 $\mathbb{P}_j[T_j > (k+1)N] =$

=

=

\leq

Now repeat k times.

Positive and null recurrence

$$(4) \quad \mathbb{E}_j[T_j] = \sum_{n=1}^{\infty} \mathbb{P}_j[T_j \geq n] =$$

$$(5) \quad \mathbb{P}_j[T_j \geq n] \text{ is } |$$

Therefore $\forall n \in \{kN+1, \dots, (k+1)N\}$

$$\mathbb{P}_j[T_j \geq n] \leq$$

$$(6) \quad \sum_{n=kN+1}^{(k+1)N} \mathbb{P}_j[T_j \geq n] \leq$$

Finally, $\mathbb{E}_j[T_j] \leq$

Conclusion: All states of an irreducible MC with finite state space are positive recurrent.

Positive recurrence and stationary distributions

Thm 9.6 Let (X_n) be a Markov chain with a state space that is countable (but not necessarily finite).

Suppose there exists a positive recurrent state $i \in S$, $\mathbb{E}_i[T_i] < \infty$.

For each state $j \in S$ define

$$\gamma(i,j) =$$

(the expected number of visits to j before reaching i).

Then the function $\pi: S \rightarrow [0,1]$

$$\pi(j) =$$

is a stationary distribution for (X_n) .

Proof.

Positive recurrence and stationary distribution

Thm 10.2 Let (X_n) be a time homogeneous MC with state space S , and suppose that the chain possesses a stationary distribution π .

- (1) If (X_n) is irreducible, then for all $j \in S$
- (2) In general, if $\pi(j) > 0$, then j is .

Proof. (1) Fix $j \in S$.

- π is stationary $\Rightarrow \pi = \pi P = \pi P^n \Leftrightarrow$
- π is distribution \Rightarrow
- (X_n) is irreducible $\Rightarrow \exists n_0 \in \mathbb{N}$ s.t.

$$\Rightarrow \pi(j) = \sum_{i \in S} \pi(i) p_{n_0}(i, j) \geq$$

Positive recurrence and stationary distribution

(2) Suppose that $\pi(j) > 0$ and j is not positive recurrent.

(i) $E_{\pi} \left[\sum_{m=1}^n \mathbb{1}_{\{X_m=j\}} \right] =$

$(E_{\pi}: \text{initial distribution is } \pi, P_{\pi}[X_0=i] = \pi(i))$

Proof: $E_{\pi} \left[\sum_{m=1}^n \mathbb{1}_{\{X_m=j\}} \right] =$

Denote $V_n(j) := \sum_{m=1}^n \mathbb{1}_{\{X_m=j\}}$, $T_j^k = \min \{n \geq 0 : V_n(j) = k\}$ - time of
k-th visit to j

(ii) $P_{\pi} \left[\lim_{K \rightarrow \infty} \frac{T_j^K}{K} = \infty \right] = 1$

Proof: If j is transient, then

(visiting j only finitely many times).

Positive recurrence and stationary distribution

Suppose that j is null recurrent. Denote

$$\tau_j^k := \text{, so that } T_j^{k+1} = \tau_j^1 + \tau_j^2 + \tau_j^3 + \dots + \tau_j^k$$

- τ_j^k are stopping times

- SMP implies that

$\{\tau_j^1, \tau_j^2, \dots\}$ are i.i.d.,

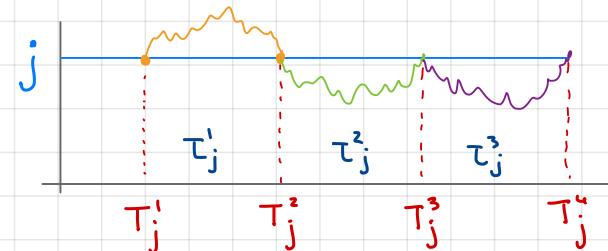
- τ_j^k have the same distribution as

$$T_j =$$

- j is null recurrent \Rightarrow

$$\frac{T_j^k}{k} = \frac{\tau_j^1 + \tau_j^2 + \dots + \tau_j^{k-1}}{k} =$$

$$\mathbb{P}_{\pi} \left[\lim_{k \rightarrow \infty} \frac{T_j^k}{k} = 0 \right] = 1, \quad \mathbb{P}_{\pi} \left[\lim_{k \rightarrow \infty} \frac{\tau_j^1 + \dots + \tau_j^{k-1}}{k-1} \cdot \frac{k-1}{k} = \infty \right] = \text{SLLN}$$



Positive recurrence and stationary distribution

(iii) $\pi(j) =$

Fix any $M > 0$.

- (ii) $\Rightarrow \exists N$ s.t.
 - $T_j^N / N \leq M$ is equivalent to $\min\{n \geq 0 : V_n(j) = N\} \leq MN$
- $V_{MN}(j) \geq N$ implies $T_j^N \leq MN$, therefore

$$\mathbb{P}_{\pi}[V_{MN}(j) \geq N] \leq$$

• $\mathbb{E}_{\pi}[V_{MN}(j)] \stackrel{(i)}{=} =$

• $\sum_{k=1}^{MN} \mathbb{P}_{\pi}[V_{MN}(j) \geq k] =$

• $MN\pi(j) < 2N \Rightarrow$

$$V_n(j) := \sum_{m=1}^n \mathbb{1}_{\{X_m=j\}}$$

Conclusion: $\pi(j) = 0$, contradiction \Rightarrow