

# MATH 285: Stochastic Processes

[math-old.ucsd.edu/~ynemish/teaching/285](http://math-old.ucsd.edu/~ynemish/teaching/285)

Today: Transition probabilities.  
Hitting times

- Test Homework on Gradescope

## Last time

Def 1.5 Let  $X_n$  be a discrete time stochastic process with state space  $S$  that is finite or countably infinite.

Then  $X_n$  is called a discrete time Markov chain if for each  $n \in \mathbb{N}$  and each  $(i_1, \dots, i_n) \in S^n$

$$(M) \quad \mathbb{P}[X_n = i_n \mid X_1 = i_1, \dots, X_{n-1} = i_{n-1}] = \mathbb{P}[X_n = i_n \mid X_{n-1} = i_{n-1}]$$

Example 1.2 (Recall  $\{X_i\}$  are i.i.d.)

Suppose that  $S$  is finite or countably infinite

Then (by independence)  $\mathbb{P}[X_n = i_n \mid X_1 = i_1, \dots, X_{n-1} = i_{n-1}] = \mathbb{P}[X_n = i_n]$  and  $\mathbb{P}[X_n = i_n \mid X_{n-1} = i_{n-1}] = \mathbb{P}[X_n = i_n]$ , so (M) is satisfied.

$$\mathbb{P}[X_1 = i_1, \dots, X_n = i_n] = \mathbb{P}[X_1 = i_1, \dots, X_n = i_n]$$

## Discrete time Markov chain

Exemple 1.2 (cont.) Recall  $S_n = X_1 + \dots + X_n$ , so  $X_n = S_n - S_{n-1}$

and thus  $\mathbb{P}[S_1 = i_1, \dots, S_n = i_n] = \mathbb{P}[X_1 = i_1, X_2 = i_2 - i_1, \dots, X_n = i_n - i_{n-1}]$

Check (M)

$$\begin{aligned}\mathbb{P}[S_n = i_n \mid S_1 = i_1, \dots, S_{n-1} = i_{n-1}] &= \frac{\mathbb{P}[S_1 = i_1, S_2 = i_2, \dots, S_n = i_n]}{\mathbb{P}[S_1 = i_1, S_2 = i_2, \dots, S_{n-1} = i_{n-1}]} \\ &= \frac{\mathbb{P}[X_1 = i_1, X_2 = i_2 - i_1, \dots, X_{n-1} = i_{n-1} - i_{n-2}, X_n = i_n - i_{n-1}]}{\mathbb{P}[X_1 = i_1, X_2 = i_2 - i_1, \dots, X_{n-1} = i_{n-1} - i_{n-2}]} \\ &= \mathbb{P}[X_n = i_n - i_{n-1}] \stackrel{\{S_{n-1} = i_{n-1}, S_{n-1} + X_n = i_n\} = \{S_{n-1} = i_{n-1}, X_n = i_n - i_{n-1}\}}{=} \\ \mathbb{P}[S_n = i_n \mid S_{n-1} = i_{n-1}] &= \frac{\mathbb{P}[S_{n-1} = i_{n-1}, S_n = i_n]}{\mathbb{P}[S_{n-1} = i_{n-1}]} = \frac{\mathbb{P}[S_{n-1} = i_{n-1}, X_n = i_n - i_{n-1}]}{\mathbb{P}[S_{n-1} = i_{n-1}]} \\ &= \mathbb{P}[X_n = i_n - i_{n-1}]\end{aligned}$$

We conclude that  $S_n$  is also a Markov chain.

## Transition probabilities. Time-homogeneous MC

"Distribution" of a Markov chain is completely described by the collection  $\{ \mathbb{P}[X_n=j | X_{n-1}=i] : (i,j) \in S^2, n \in \mathbb{N} \}$

Def. 1.6 A Markov chain is called **time-homogeneous** if for any  $i, j \in S$   $\mathbb{P}[X_n=j | X_{n-1}=i]$  is independent of  $n$  i.e., there exists a function  $p: S \times S \rightarrow [0,1]$  s.t.

$$\mathbb{P}[X_n=j | X_{n-1}=i] = p(i,j) \quad \forall n \in \mathbb{N}$$

We call  $\mathbb{P}[X_n=j | X_{n-1}=i]$  the transition probabilities

"Distribution" of a time-homogeneous MC is determined by the transition probabilities  $p(i,j)$  and the distribution of  $X_1$  (initial distribution)

# Transition probabilities

If  $p(i,j)$  are the transition probabilities, then

$$\sum_{j \in S} p(i,j) = \sum_{j \in S} \mathbb{P}[X_2=j | X_1=i] = \sum_{j \in S} \frac{\mathbb{P}[X_2=j, X_1=i]}{\mathbb{P}[X_1=i]} = \frac{\mathbb{P}[X_1=i]}{\mathbb{P}[X_1=i]} = 1$$

Def. If  $A$  is an  $n \times n$  matrix s.t.  $\forall i \in \{1, \dots, n\}$

$\sum_{j=1}^n A_{ij} = 1$ , then  $A$  is called stochastic

Suppose  $|S| < \infty$  and let

$$P = [p(i,j)]_{i,j \in S}$$

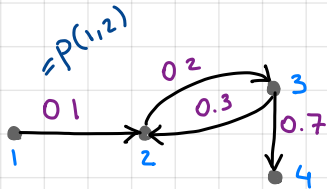
$$P = \begin{matrix} & \begin{matrix} s_1 & s_2 & \dots \end{matrix} \\ \begin{matrix} s_1 \\ s_2 \\ \vdots \end{matrix} & \begin{bmatrix} p(s_1, s_1) & p(s_1, s_2) & \dots \\ p(s_2, s_1) & p(s_2, s_2) & \dots \\ \vdots & \vdots & \ddots \end{bmatrix} \end{matrix}$$

↑  
transition matrix

Then  $P$  is stochastic

Ex.

$$S = \{1, 2, 3, 4\}$$



$$P = [p(i,j)] = \begin{matrix} & \begin{matrix} 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \begin{bmatrix} 0.9 & 0.1 & 0 & 0 \\ 0 & 0.8 & 0.2 & 0 \\ 0 & 0.3 & 0 & 0.7 \\ 0 & 0 & 0 & 1 \end{bmatrix} \end{matrix}$$

# Transition probabilities

## Example 1.7

Markov chain on  $S = \{0, 1, 2, \dots, N\}$



Transition probabilities:

if  $i \in \{1, 2, \dots, N-1\}$  then

$$p(i, j) = \begin{cases} \frac{1}{2}, & \text{if } j = i+1 \\ \frac{1}{2}, & \text{if } j = i-1 \\ 0, & \text{otherwise} \end{cases}$$

Reflecting random walk:

$$p(0, 1) = p(N, N-1) = 1$$

Absorbing random walk:

$$p(0, 0) = p(N, N) = 1$$

Partially reflecting walk:

$$p(0, 0) = p \quad p(0, 1) = 1-p$$

$$p(N, N) = q \quad p(N, N-1) = 1-q$$

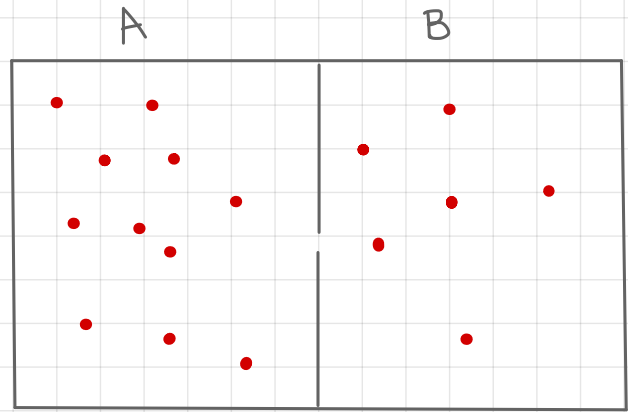
# Ehrenfest model

Diffusion through a membrane:

$N$  particles in two chambers

connected by a small hole

[Think about two urns with  
 $N$  balls]



At each step you choose (unif. at random) a ball and put it to the other urn [particle passing through the membrane].

Denote  $X_n = \#$  balls in urn A.  $X_n \in \{0, 1, \dots, N\}$

$(X_n)$  is a Markov chain

$$\forall i \in \{0, 1, \dots, N\} \quad p(i, i+1) = \frac{i}{N} \cdot \frac{1}{2}$$
$$p(i, i-1) = \frac{i-1}{N} \cdot \frac{1}{2}$$

$$P = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & \dots & N \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ \dots \end{matrix} & \begin{bmatrix} 0 & 1 & 0 & 0 & \dots & 0 \\ 0 & -1/2 & 1/2 & 0 & \dots & 0 \\ 0 & 0 & -1/2 & 1/2 & \dots & 0 \\ 0 & 0 & 0 & -1/2 & 1/2 & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & 0 & \dots & 1 \end{bmatrix} \end{matrix}$$

## n-step transition probabilities

Let  $(X_n)_{n \in \mathbb{N} \cup \{0\}}$  be a Markov chain (time-homogeneous, finite  $S$ )

**Q:** Given  $X_0$ , what is the distribution of  $X_n$ ?

Lemma 2.3 Let  $(X_n)$  be a time-homogeneous Markov chain with a finite state space  $S$  and transition matrix  $P$ .

Then  $\forall n \in \mathbb{N}$   $\mathbb{P}[X_n = j \mid X_0 = i] = [P^n]_{ij}$

Proof (Induction by  $n$ )  $n=1$ :  $\mathbb{P}[X_1 = j \mid X_0 = i] = P_{ij}$

Induction step: Suppose that  $\mathbb{P}[X_n = j \mid X_0 = i] = [P^n]_{ij}$ . Then

$$\begin{aligned}\mathbb{P}[X_{n+1} = j \mid X_0 = i] &= \sum_{k \in S} \mathbb{P}[X_{n+1} = j, X_n = k \mid X_0 = i] \\ &= \sum_{k \in S} \mathbb{P}[X_{n+1} = j \mid X_n = k, X_0 = i] \mathbb{P}[X_n = k \mid X_0 = i] \\ &= \sum_{k \in S} \mathbb{P}[X_{n+1} = j \mid X_n = k] [P^n]_{ik} = \sum_k [P^n]_{ik} P_{kj} \\ &= [P^{n+1}]_{ij} \quad \blacksquare\end{aligned}$$



# Chapman - Kolmogorov Equations

Cor. 2.4 Let  $(X_n)$  be as in Lem 2.3. Denote by  $p_n(i, j)$  the  $n$ -step transition probability  $p_n(i, j) = \mathbb{P}[X_n = j \mid X_0 = i]$ .

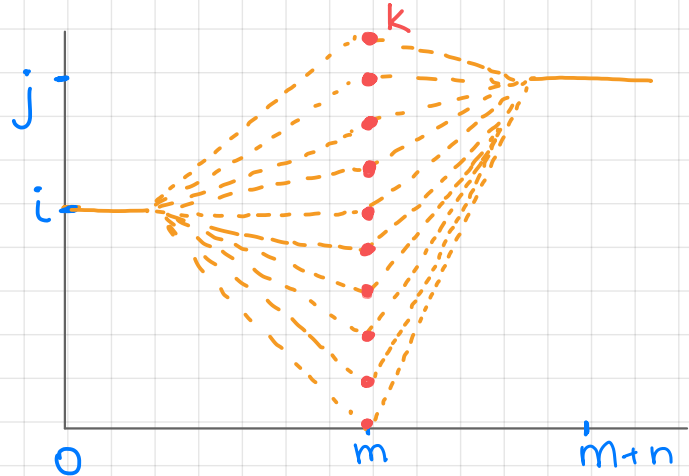
Then for any  $m, n \in \mathbb{N}$

$$p_{m+n}(i, j) = \sum_{k \in S} p_m(i, k) p_n(k, j)$$

Proof. By Lem 2.3,  $\forall n \in \mathbb{N}$   $p_n(i, j) = [P^n]_{ij}$ . Then

$$p_{m+n}(i, j) = [P^{m+n}]_{ij} = \sum_k P_{ik}^m P_{kj}^n$$

$$= \sum_k p_m(i, k) p_n(k, j) \quad \blacksquare$$



## Markov property "future is independent of the past"

Prop 2.5 Let  $(X_n)$  be a time-homogeneous MC with discrete state space  $S$  and transition probabilities  $p(i, j)$ . Fix  $m \in \mathbb{N}$ ,  $l \in S$ , and suppose that  $\mathbb{P}[X_m = l] > 0$ . Then **conditional on  $X_m = l$** , the process  $(X_{m+n})_{n \in \mathbb{N}}$  is **Markov** with transition probabilities  $p(i, j)$  initial distribution  $\delta_l$  (atom at  $l$ ) and **independent** of the random variables  $X_0, X_1, \dots, X_m$ .

Proof. Let  $A$  be an event determined by  $X_0, \dots, X_m$ .

- First assume that  $(i_0, \dots, i_m) \in S^{m+1}$  for some  $(i_0, \dots, i_m) \in S^{m+1}$

Then  $\mathbb{P}[\{X_m = i_m, \dots, X_{m+n} = i_{m+n}\} \cap A \mid X_m = l]$

=

