

# MATH 285: Stochastic Processes

[math-old.ucsd.edu/~ynemish/teaching/285](http://math-old.ucsd.edu/~ynemish/teaching/285)

## Today: Hitting times. First step analysis

- Test Homework on Gradescope

## Initial distribution and transition matrix

Let  $(X_n)_{n \geq 0}$  be a (time-homogeneous) Markov chain with finite state space  $S = \{s_1, s_2, \dots, s_{|S|}\} (= \{1, 2, 3, \dots, |S|\})$

Distribution of  $X_n$  is a vector  $(\mathbb{P}[X_n=1], \mathbb{P}[X_n=2], \dots, \mathbb{P}[X_n=|S|])$

Let  $\lambda = (\lambda_1, \lambda_2, \dots, \lambda_{|S|})$  be the distribution of  $X_0$ , i.e.,  $\mathbb{P}[X_0=i] = \lambda_i$ . Let  $P$  be the transition matrix of  $(X_n)$ .

Q: What is the distribution of  $X_n$ ?

$$X_1: \mathbb{P}[X_1=j] = \sum_{i \in S} \mathbb{P}[X_1=j | X_0=i] \mathbb{P}[X_0=i] = \sum_{i=1}^{|S|} \lambda_i p^{(i,j)} = [\lambda P]_j$$

Distribution of  $X_1$  is given by  $\lambda P$

$$X_n: \mathbb{P}[X_n=j] = \sum_{i=1}^{|S|} \mathbb{P}[X_n=j | X_0=i] \mathbb{P}[X_0=i] = \sum_{i=1}^{|S|} \lambda_i p_n^{(i,j)} = [\lambda P^n]_j$$

Distribution of  $X_n$  is given by  $\lambda P^n$

We will say that  $(X_n)$  is Markov  $(\lambda, P)$

## Markov property "future is independent of the past"

Prop 2.5 Let  $(X_n)$  be a time-homogeneous MC with discrete state space  $S$  and transition probabilities  $p(i, j)$ . Fix  $m \in \mathbb{N}$ ,  $l \in S$ , and suppose that  $\mathbb{P}[X_m = l] > 0$ . Then **conditional on  $X_m = l$** , the process  $(X_{m+n})_{n \in \mathbb{N}}$  is Markov with transition probabilities  $p(i, j)$ , initial distribution  $(0, \dots, 0, \overset{l-1}{\uparrow} 1, \overset{l}{\downarrow} 0, \dots, 0)$  and **independent** of the random variables  $X_0, \dots, X_m$ , i.e. if  $A$  is an event determined by  $X_0, X_1, \dots, X_m$  and  $\mathbb{P}[A \cap \{X_m = l\}] > 0$  then for all  $n \geq 0$

$$\mathbb{P}[X_{m+1} = i_{m+1}, \dots, X_{m+n} = i_{m+n} \mid A \cap \{X_m = l\}] = p(l, i_{m+1}) p(i_{m+1}, i_{m+2}) \dots p(i_{m+n-1}, i_{m+n})$$

Proof. Enough to show that (\*)

$$\mathbb{P}[\{X_{m+1} = i_{m+1}, \dots, X_{m+n} = i_{m+n}, X_m = l\} \cap A] = p(l, i_{m+1}) \dots p(i_{m+n-1}, i_{m+n}) \mathbb{P}[A \cap \{X_m = l\}]$$

# Markov property

- Let  $A = \{X_0 = i_0, \dots, X_m = l\}$ . Then

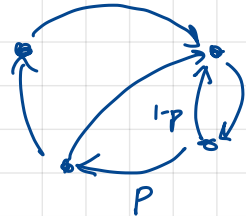
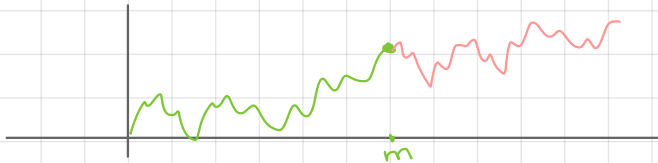
$$\mathbb{P}[X_0 = i_0, \dots, X_m = l, X_{m+1} = i_{m+1}, \dots, X_{m+n} = i_{m+n}] = \mathbb{P}[X_0 = i_0] p(i_0, i_1) \dots p(i_{m-1}, l) \times p(l, i_{m+1}) \dots p(i_{m+n-1}, i_{m+n})$$

$$\mathbb{P}[X_0 = i_0, \dots, X_m = l] = \mathbb{P}[X_0 = i_0] p(i_0, i_1) \dots p(i_{m-1}, l)$$

- Any set  $A$  determined by  $X_0, \dots, X_m$  is a disjoint union of the events of the form  $\{X_0 = i_0, \dots, X_m = i_m\}$ .

$$\begin{aligned} \text{E.g. } \mathbb{P}[\{X_{m+1} = i_{m+1}, \dots, X_{m+n} = i_{m+n}\} \cap (A_1 \cup A_2) \cap \{X_m = l\}] \\ = p(l, i_{m+1}) \dots p(i_{m+n-1}, i_{m+n}) (\mathbb{P}[A_1 \cap \{X_m = l\}] + \mathbb{P}[A_2 \cap \{X_m = l\}]) \\ = p(l, i_{m+1}) \dots p(i_{m+n-1}, i_{m+n}) \mathbb{P}[(A_1 \cup A_2) \cap \{X_m = l\}] \end{aligned}$$

So (\*) holds for any event  $A$ .



# Hitting times

Q1: When is the first time the process enters a certain set?

For  $A \subset S$ , compute  $\tau_A := \min \{n \in \mathbb{N} \cup \{0\} : X_n \in A\}$

Q2: For  $A, B \subset S$ ,  $A \cap B = \emptyset$  find the probability

$$\mathbb{P}[\tau_A < \tau_B \mid X_0 = i] : h(i)$$

Start with Q2

• trivial: 
$$h(i) = \begin{cases} 1, & i \in A \\ 0, & i \in B \end{cases}$$

• take  $i \notin A \cup B$ ; "first step analysis":

$$\mathbb{P}[\tau_A < \tau_B \mid X_0 = i] = \sum_{j \in S} \mathbb{P}[\tau_A < \tau_B \mid X_1 = j, X_0 = i] \mathbb{P}(X_1 = j \mid X_0 = i)$$

By the Markov property

$$\mathbb{P}[\tau_A < \tau_B \mid X_0 = i, X_1 = j] = \mathbb{P}[\tau_A < \tau_B \mid X_1 = j] = \mathbb{P}[\tau_{A-1} < \tau_{B-1} \mid X_0 = j] = h(j)$$

$X_k \in A \Rightarrow \tau_A = k$

## Hitting times

We conclude that

$$h(i) = \sum_{j \in S} p(i,j) h(j) \quad (**)$$

This gives a system of linear equations + boundary conditions

$$h(i) = \begin{cases} 1, & i \in A \\ 0, & i \in B \end{cases} \quad (***)$$

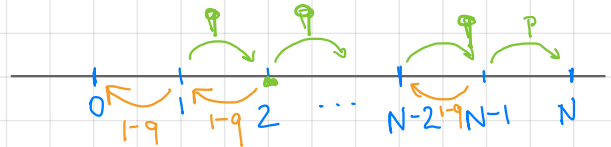
If  $S$  is finite, denote  $\bar{h} := (h(1), h(2), \dots, h(|S|))$ . Then

(\*\*) becomes  $\bar{h} = Ph$

Example 2.6  $(X_n)$  random walk on  $\{0, 1, 2, \dots, N\}$ , not necessarily symmetric,  $p(i, i+1) = q$ ,  $p(i, i-1) = 1-q$ ,  $q \in [0, 1]$

Let  $i \in \{1, 2, \dots, N-1\}$ . Compute

$\mathbb{P}[X_n \text{ reaches } N \text{ before } 0 \mid X_0 = i]$



# Hitting times for random walks

Denote  $A = \{N\}$ ,  $B = \{0\}$ . Need  $\mathbb{P}[\tau_A < \tau_B \mid X_0 = i] = h(i)$

- boundary conditions  $h(0) = 0$ ,  $h(N) = 1$

Consider  $0 < i < N$

- recall  $p(i, j) = \begin{cases} q, & j = i+1 \\ 1-q, & j = i-1 \\ 0, & \text{otherwise} \end{cases}$ , so  $(**)$  becomes  $h(i) = \sum_{j \in S} p(i, j) h(j)$

$$h(i) = q h(i+1) + (1-q) h(i-1)$$

$$(1-q)(h(i) - h(i-1)) = q(h(i+1) - h(i))$$

$$\forall i \in \{1, \dots, N-1\}$$

• if  $q = 0$ , then  $h(i) = h(i-1) = h(0) = 0$

• if  $q = 1$ , then  $h(i) = 1$

• if  $q \in (0, 1)$ , denote  $\Delta h(i) := h(i) - h(i-1)$ ,  $\theta := \frac{1-q}{q}$

# Hitting times for random walks

$$\textcircled{+} \begin{cases} \Delta h(1) = \Delta h(1) \\ \Delta h(2) = \theta \Delta h(1) \\ \Delta h(3) = \theta \Delta h(2) = \theta^2 \Delta h(1) \\ \vdots \\ \Delta h(N) = \theta^{N-1} \Delta h(1) \end{cases}$$

Take the sum of the first  $i$  equations

$$\text{LHS: } \Delta h(1) + \Delta h(2) + \dots + \Delta h(i) = \cancel{h(1)} - h(0) + h(2) - \cancel{h(1)} \dots = h(i) - h(0)$$

$$\text{RHS: } (1 + \theta + \theta^2 + \dots + \theta^{i-1}) \Delta h(1)$$

$$\Rightarrow \forall i \in \{2, 3, \dots, N\} \quad h(i) = \sum_{\ell=0}^{i-1} \theta^\ell \Delta h(1)$$

$$h(N) = 1 = \sum_{\ell=0}^{N-1} \theta^\ell \Delta h(1) \quad \Rightarrow \Delta h(1) = h(1) = \frac{1}{\sum_{\ell=0}^{N-1} \theta^\ell}$$

$$\Rightarrow h(i) = \frac{\sum_{\ell=0}^{i-1} \theta^\ell}{\sum_{\ell=0}^{N-1} \theta^\ell}$$

for  $i \in \{1, \dots, N-1\}$



## Gambler's ruin

Suppose you have 100\$, at each game you bet 1\$, and you stop either when your fortune reaches 200\$ or when you lose everything. [  $N=200$ ,  $h(100)=?$  ]

(fair game) If probability of winning is 0.5 ( $q=0.5$ )  
then  $\theta = \frac{0.5}{0.5} = 1$ ,  $h(100) = \frac{100}{200} = \frac{1}{2} = 0.5$

(real gambling) If probability of winning is  $\frac{18}{38}$  ( $q=0.474$ )  
then  $h(100) = \frac{1 - \theta^{100}}{1 - \theta^{200}} = 0.000027$