

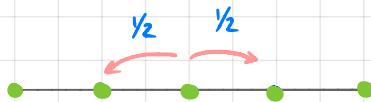
MATH 285: Stochastic Processes

math-old.ucsd.edu/~ynemish/teaching/285

Today: Stationary distribution

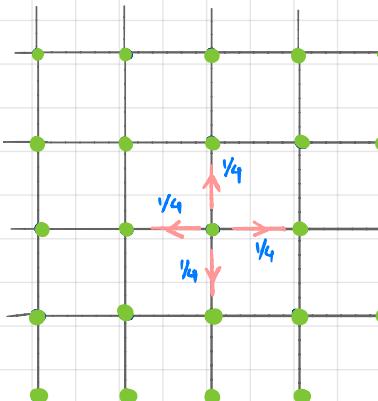
- Homework 1 is due on Friday, January 14, 11:59 PM

SSRW on \mathbb{Z}^d , $d \in \{1, 2, 3\}$

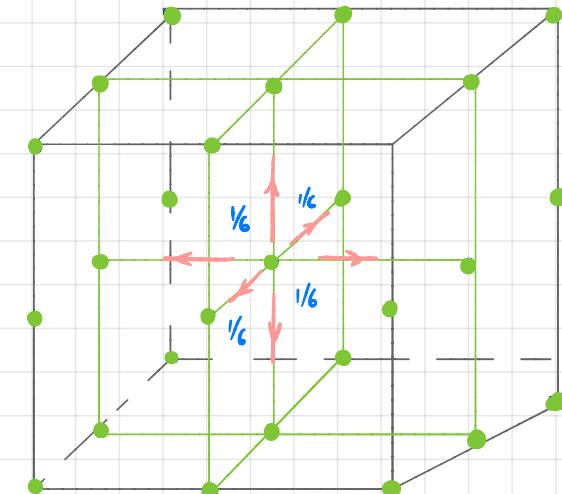


~~transient~~

recurrent

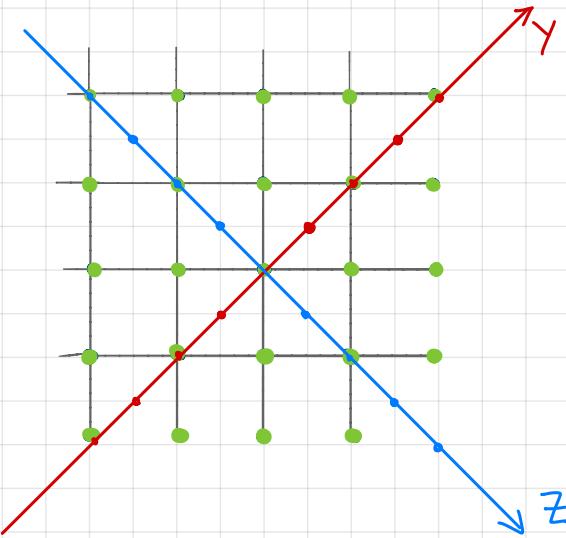


transient
recurrent



transient
~~recurrent~~

Simple symmetric random walk on \mathbb{Z}^2



Y_n = projection of X_n on $y=x$

Z_n = projection of X_n on $y=-x$

$X_n = (i, j) \Leftrightarrow Y_n = i+j, Z_n = j-i$

$X_n = (0, 0) \Leftrightarrow Y_n = 0, Z_n = 0$

Let (Y_n) and (Z_n) be two independent SSRW on \mathbb{Z}

Define $\tilde{X}_n = \left(\frac{Y_n + Z_n}{2}, \frac{Y_n - Z_n}{2} \right)$

Then (\tilde{X}_n) is a SSRW on \mathbb{Z}^2 ; $\tilde{X}_n = (\bar{0}, \bar{0}) \Leftrightarrow Y_n = 0, Z_n = 0$

$$P_n^{\tilde{x}}(\bar{0}, \bar{0}) = \mathbb{P}[\tilde{X}_n = (0, 0) \mid \tilde{X}_0 = (0, 0)] = \mathbb{P}[Y_n = 0, Z_n = 0 \mid Y_0 = 0, Z_n = 0]$$

$$= P_n^Y(0, 0) P_n^Z(0, 0) \sim \left(\frac{1}{\sqrt{\pi n}}\right)^2 = \frac{1}{\pi n} \Rightarrow \sum_{n=0}^{\infty} P_n^{\tilde{x}}(\bar{0}, \bar{0}) \sim \sum_{n=0}^{\infty} \frac{1}{\pi n} = \infty$$

Markov processes

Let (X_n) be a Markov chain with initial distribution λ and transition matrix P .

- Distribution of X_n : λP^n
- First step analysis:
 - absorption probabilities (gambler's ruin)
 - mean hitting times (two consecutive heads)
- Class structure : recurrence / transience
 - criteria
 - SSRW on $\mathbb{Z}, \mathbb{Z}^2, \mathbb{Z}^3$
- Irreducibility

Long-run behavior of Markov chains

Denote by π_n the distribution of X_n , i.e.,

$$\pi_n = (\mathbb{P}[X_n=1], \mathbb{P}[X_n=2], \dots, \mathbb{P}[X_n=|S|])$$

$$\pi_n = \pi_0 P^n \quad (\text{follows from the Chapman-Kolmogorov eqs.})$$

What happens with π_n as $n \rightarrow \infty$?

P^n as $n \rightarrow \infty$ for a stochastic matrix P

Examples : ② $P_2 = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$ ③ $P^n = \begin{bmatrix} 0.04 & 0.05 & 0.91 \\ 0.04 & 0.05 & 0.91 \\ 0.04 & 0.05 & 0.91 \end{bmatrix}$

① $P_1 = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ $P_2^{2n} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ $\pi_n = \pi_0 P_n$

$P_1^n = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$ $P_2^{2n+1} = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$ $\pi_n = (\pi_0(1) \cdot 0.04 + \pi_0(2) \cdot 0.04 + \pi_0(3) \cdot 0.04)$
 $= 0.04, 0.05, 0.91$

$$\pi_n = \pi_0$$

$$\pi_{2n} = \pi_0, \pi_{2n+1} = (\pi_0(2), \pi_0(1))$$

Stationary distribution

Def 6.1 Let $(X_n)_{n \geq 0}$ be a Markov chain with state space S and transition matrix P . A vector $\pi = (\pi(i))_{i \in S}$ is called a stationary distribution if $\pi(i) \geq 0$ for all $i \in S$, $\sum_{i \in S} \pi(i) = 1$ and $\pi P = \pi$ (*)

If π is the stationary distribution and $\pi_0 = \pi$, then

$$\forall n \quad \pi_n = \pi P^n = \pi P \cdot P^{n-1} = \pi \cdot P^{n-1} = \dots = \pi$$

In order to find the stationary distribution we have to solve the linear system (*):

- π is the left eigenvector of P with e.v. 1

Stationary distribution

Q 1: Existence of the stationary distribution

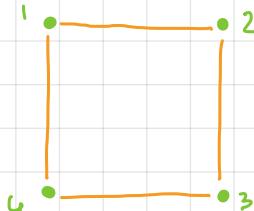
Q 2: Uniqueness of the stationary distribution

Q 3: Convergence to the stationary distribution

Examples 6.3. (1) $S = \mathbb{Z}$, $p(i, i+1) = 1 \quad \forall i \in \mathbb{Z}$ (deterministic).

Then $\forall i \lim_{n \rightarrow \infty} \mathbb{P}[X_n = i] = 0$, so st. distr. does not exist

(2) $S = \{1, 2, 3, 4\}$, $P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ 0 & 0 & \frac{1}{2} & \frac{1}{2} \\ 0 & 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix}$. Then $\pi = (\frac{1}{2}, \frac{1}{2}, 0, 0)$ and $\pi' = (0, 0, \frac{1}{2}, \frac{1}{2})$ are both stationary distributions.

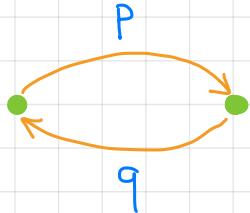
(3) SSRW on 
 $P[X_{2n+1} \in \{1, 3\}] = 0$
 $P[X_{2n} \in \{1, 3\}] = 1$

If $X_0 = 1$, then $P[X_{2n+1} \in \{1, 3\}] = 0$
 $P[X_{2n} \in \{1, 3\}] = 1$

$P = \begin{bmatrix} 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} & 0 \\ 0 & \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} & 0 \end{bmatrix}$

$\pi = (\frac{1}{4}, \frac{1}{4}, \frac{1}{4}, \frac{1}{4})$

General 2-state Markov chain



$$P = \begin{bmatrix} 1-p & p \\ q & 1-q \end{bmatrix} \quad p, q \in [0, 1]$$

$$\det(P - \lambda I) = (1-p-\lambda)(1-q-\lambda) - pq = \lambda^2 + \lambda(p+q-2) + 1-p-q = 0$$

↳ eigenvalues are $1, 1-p-q$

$$P - \lambda I$$

$$P - I = \begin{pmatrix} -p & p \\ q & -q \end{pmatrix} \quad \left\{ \begin{array}{l} -\pi(1)p + \pi(2) \cdot q = 0 \\ \pi(1) + \pi(2) = 1 \end{array} \right. \quad P - (1-p-q)I = \begin{bmatrix} q & p \\ q & p \end{bmatrix}$$

$$P \in \left\{ \begin{bmatrix} 1 & 0 \\ 1 & 0 \end{bmatrix}, \begin{bmatrix} 0 & 1 \\ 0 & 1 \end{bmatrix}, \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}, \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \right\}$$

Case 1:

$$p, q \in \{0, 1\}$$

$$\pi = (1, 0)$$

unique

$$\pi = (0, 1)$$

unique

$$\text{any } \pi$$

$$\pi(1) + \pi(2) = 1$$

not unique

$$\pi = \left(\frac{1}{2}, \frac{1}{2}\right)$$

unique

$$P^n = P$$

$$P^n = \begin{bmatrix} 0 & 1 \\ 0 & 1 \end{bmatrix}$$

$$P^n = P$$

$$P^{2n+1} = P, P^{2n} = I$$

$$\text{no convergence}$$

General 2-state Markov chain

Case 2: $p \in (0, 1)$ or $q \in (0, 1)$

$$\begin{cases} -\pi(1)p + \pi(2)q = 0 \\ \pi(1) + \pi(2) = 1 \end{cases} \Rightarrow \pi(1) = \frac{q}{p+q}, \pi(2) = \frac{p}{p+q}$$

$$(x, y) \begin{pmatrix} q & p \\ q & p \end{pmatrix} = (0, 0) \Rightarrow x = -y \quad (x, y) = (1, -1)$$

$$Q^{-1} = \begin{pmatrix} \frac{q}{p+q} & \frac{p}{p+q} \\ 1 & -1 \end{pmatrix}, \quad Q = -\begin{bmatrix} -1 & -\frac{p}{p+q} \\ -1 & \frac{q}{p+q} \end{bmatrix} = \begin{bmatrix} 1 & \frac{p}{p+q} \\ 1 & -\frac{q}{p+q} \end{bmatrix}$$

$$P = Q \begin{bmatrix} 1 & 0 \\ 0 & 1-p-q \end{bmatrix} Q^{-1} \Rightarrow P^n = Q \begin{bmatrix} 1 & 0 \\ 0 & (1-p-q)^n \end{bmatrix} Q^{-1} \xrightarrow{n \rightarrow \infty} Q \begin{bmatrix} 1 & 0 \\ 0 & 0 \end{bmatrix} Q^{-1}$$

$$\lim_{n \rightarrow \infty} P^n = \begin{bmatrix} 1 & \frac{p}{p+q} \\ 1 & -\frac{q}{p+q} \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & 0 \end{bmatrix} \begin{bmatrix} \frac{q}{p+q} & \frac{p}{p+q} \\ 1 & -1 \end{bmatrix} = \begin{bmatrix} 1 \\ 1 \end{bmatrix} \begin{bmatrix} \frac{q}{p+q} & \frac{p}{p+q} \end{bmatrix} = \begin{bmatrix} \pi \\ \pi \end{bmatrix}$$

$\lim_{n \rightarrow \infty} \pi_n = \pi$ regardless of initial distribution.

General Markov chain with finite state space

Let (X_n) be a MC with finite state space S .

Suppose that $\pi = P\pi$, $P = QDQ^{-1}$ such that

$$Q = \begin{bmatrix} 1 & & \\ 1 & * & \\ \vdots & & \\ 1 & & \end{bmatrix}, \quad Q^{-1} = \begin{bmatrix} \pi & & \\ & * & \\ & & \end{bmatrix}, \quad D = \begin{bmatrix} 1 & 0 & \cdots & 0 \\ 0 & M & & \\ \vdots & & & \\ 0 & & & \end{bmatrix}, \quad \lim_{n \rightarrow \infty} M^n = 0.$$

$$\text{Then } \lim_{n \rightarrow \infty} P^n = \lim_{n \rightarrow \infty} Q D^n Q^{-1} = \begin{bmatrix} 1 & & \\ \vdots & * & \\ 1 & & \end{bmatrix} \begin{bmatrix} 1 & 0 & \cdots & 0 \\ 0 & M & & \\ \vdots & & & \\ 0 & & & \end{bmatrix} \begin{bmatrix} \pi & & \\ & * & \\ & & \end{bmatrix} = \begin{bmatrix} \pi & & \\ \pi & \ddots & \\ \vdots & & \pi \end{bmatrix}$$