

MATH 285: Stochastic Processes

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Today: Reducible Markov chains with
finite state space
Markov chains with infinite state
space

- Homework 2 is due on Friday, January 21 11:59 PM

General form of transition matrix with finite S

$$P = \begin{bmatrix} P_e & & \\ & \ddots & \\ & & P_e \\ \hline & & \\ & S & Q \end{bmatrix}$$

P_e submatrix for the recurrent class R_e

P_e is a stochastic matrix,
we can consider it as a
Markov chain on R_e

[S|Q] transition probabilities starting from transient states.

- If P_e is aperiodic, then $P_e^n \rightarrow \begin{bmatrix} \pi^{(e)} \\ \vdots \\ \pi^{(e)} \end{bmatrix}$, $n \rightarrow \infty$
- What about transient states?
- What if P_e is not aperiodic?

Transient states

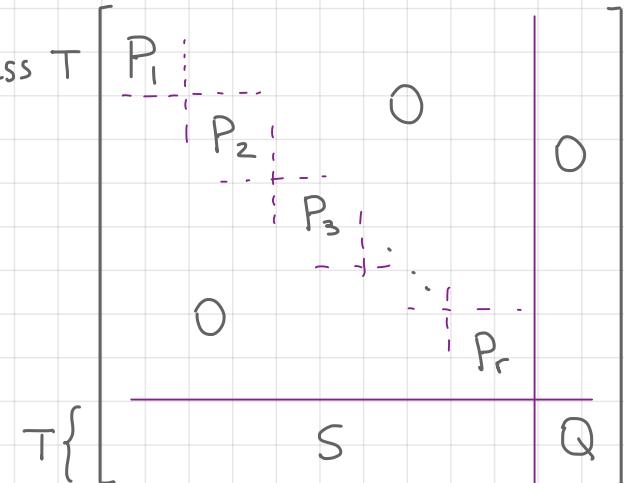
Suppose there exists one transient class T

- $S \neq 0$

If $S = 0$ then T is recurrent

- If $S \neq 0$, then Q is substochastic,

i.e., $\exists i \in T$ s.t. $\sum_{j \in T} Q_{ij} < 1$



- If Q is substochastic, then for all eigenvalues λ of Q $|\lambda| < 1$

$\Rightarrow Q^n \rightarrow 0$, $n \rightarrow \infty$, i.e. for $i, j \in T$ $P_i[X_n=j] \rightarrow 0$, $n \rightarrow \infty$

- $I + Q + Q^2 + \dots = I + \sqrt{D} \sqrt{V} + \sqrt{D}^2 \sqrt{V} + \dots = \sqrt{(I + D + D^2 + \dots)} \sqrt{V}$ converges

$$\text{For } i, j \in T, E_i \left[\sum_{k=0}^{\infty} \mathbb{1}_{X_k=j} \right] = \sum_{k=0}^{\infty} P_i[X_k=j] = p_0(i,j) + p_1(i,j) + p_2(i,j) + \dots$$

$$= S_{ij} + Q_{ij} + [Q^2]_{ij} + \dots = [(I + Q + Q^2 + \dots)]_{ij} = [(I - Q)^{-1}]_{ij}$$

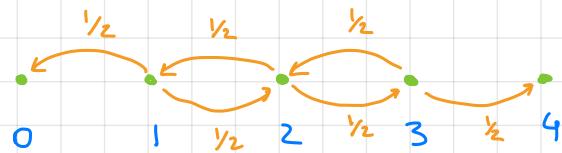
Transient states

Conclusion: if TCS is a transient class, then $\forall i, j \in T$

$$\lim_{n \rightarrow \infty} P_i[X_n = j] = 0$$

$$E_i \left[\sum_{k=0}^n \mathbb{1}_{\{X_k = j\}} \right] = [(I - Q)^{-1}]_{ij} \quad \text{expected number of visits to } j \text{ starting from } i$$

Example 8.1



$$Q = \begin{bmatrix} 0 & \frac{1}{2} & 0 \\ \frac{1}{2} & 0 & \frac{1}{2} \\ 0 & \frac{1}{2} & 0 \end{bmatrix}$$

$$(I - Q)^{-1} = \begin{bmatrix} \frac{3}{2} & 1 & \frac{1}{2} \\ 1 & 2 & 1 \\ \frac{1}{2} & 1 & \frac{3}{2} \end{bmatrix}$$

$$P = \left[\begin{array}{cc|cc} 0 & 4 & 1 & 2 & 3 \\ 1 & 0 & 0 & 0 & 0 \\ 2 & \frac{1}{2} & 0 & \frac{1}{2} & 0 \\ 3 & 0 & 0 & \frac{1}{2} & 0 \\ 4 & 0 & \frac{1}{2} & 0 & \frac{1}{2} \end{array} \right]$$

Expected number of visits to ② starting from ① is 1

Expected number of steps before absorption starting from ① is $\frac{3}{2} + 1 + \frac{1}{2} = 3$

Transient states

Recall, First step analysis for the mean hitting time

$$g_i = \mathbb{E}_i [\tau_A] = \begin{cases} 0, & i \in A \\ 1 + \sum_{j \in S} P(i,j) g_j, & i \notin A \end{cases}$$

$$\tau_A = \sum_{n=0}^{\infty} \mathbb{1}_{\{X_n \notin A\}}$$

Instead of adding 1 for each step, add 1 only when X_n visits j :

Denote $S \setminus A =: T$, and for $i, j \in T$ $g_{ij} = \mathbb{E}_i \left[\sum_{n=0}^{\infty} \mathbb{1}_{\{X_n=j\}} \right]$

Then FSA $g_{kj} = 0 \quad \text{if } k \in A$

$$g_{ij} = \delta_{ij} + \sum_{k \in S} P(i,k) g_{kj} = \delta_{ij} + \sum_{k \in T} P(i,k) g_{kj}$$

$$G = [g_{ij}] \text{, then } G = I + QG \Rightarrow G = (I - Q)^{-1}$$

Transient states

Starting from T_i , in which class will (X_n) end up?

Collapse each R_e into one state r_e ,

keep transient states t_e , $T = \{t_e\}$

(\tilde{X}_n) new MC on the reduced state

space, and transition matrix \tilde{P}_1 ,

with $\tilde{s}(t_i, r_j) = P_{t_i} [X_1 \in R_j]$

Denote $\tilde{A} = [\alpha(t_i, r_j)]$ with

$\alpha(t_i, r_j) = P_{t_i} [(X_n) \text{ enters } r_j \text{ eventually}]$

Then

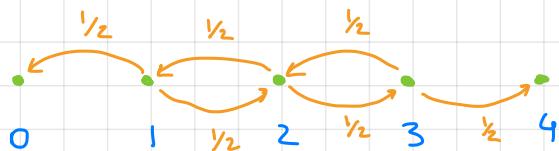
$$\tilde{A} = (I - Q)^{-1} \tilde{s}$$

$$R_1 \left\{ \begin{array}{c} P_1 \\ P_2 \\ \vdots \\ P_r \end{array} \right. \quad \left. \begin{array}{c} 0 \\ 0 \\ \ddots \\ \ddots \end{array} \right| \quad T \quad S \quad Q$$

$$\tilde{P} = \left[\begin{array}{cccccc|c} t_1 & 0 & 0 & \cdots & 0 & 0 \\ t_2 & 0 & 1 & 0 & \cdots & 0 & 0 \\ t_3 & 0 & 0 & 1 & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & \cdots & 1 & 0 & 0 \end{array} \right] \quad \tilde{s} \quad Q$$

Transient states

Example 8.2



$$P = \left[\begin{array}{cc|cc} 0 & 1 & 0 & 0 \\ 4 & 0 & 1 & 0 \\ 1 & \frac{1}{2} & 0 & \frac{1}{2} \\ 2 & 0 & 0 & \frac{1}{2} \\ 3 & 0 & \frac{1}{2} & 0 \end{array} \right] \quad Q$$

\tilde{s}

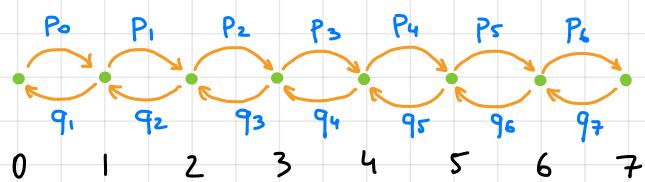
What is the probability that starting from a transient state i we end up in a recurrent state j ?

Use $\tilde{A} = (I - Q)^{-1} \tilde{s}$ (nothing to collapse in this case)

$$\tilde{A} = \left[\begin{array}{ccc} \frac{3}{2} & 1 & \frac{1}{2} \\ 1 & 2 & 1 \\ \frac{1}{2} & 1 & \frac{3}{2} \end{array} \right] \left[\begin{array}{cc} \frac{1}{2} & 0 \\ 0 & 0 \\ 0 & \frac{1}{2} \end{array} \right] = \left[\begin{array}{cc} \frac{3}{4} & \frac{1}{4} \\ \frac{1}{2} & \frac{1}{2} \\ \frac{1}{4} & \frac{3}{4} \end{array} \right]$$

- Expected transit times from i to j (think about j as absorbing) ...

Birth and death processes (infinite state space)



$$S = \{0, 1, 2, 3, \dots\}$$

$$p(i, i+1) = p_i, \quad p(i, i-1) = 1 - p_i$$

$$p(0, 1) = p_0, \quad p(0, 0) = 1 - p_0$$

$p_0 \in [0, 1]$, $p_0 = 0$ absorbing, $p_0 = 1$ reflecting

Model of population growth : X_n = size of the population at time n

$P_i[\exists_{n \geq 0} : X_n = 0]$ – extinction probability

$P_i[X_n \rightarrow \infty \text{ as } n \rightarrow \infty]$ – probability that population explodes

Denote $h(i) := P_i[\exists_{n \geq 0} X_n = 0] = P_i[\tau_0 < \infty]$, $\tau_0 = \min\{n \geq 0, X_n = 0\}$

First step analysis:

Theorem 7.0 $(h(0), h(1), \dots)$ is the minimal solution to

$$\begin{cases} h(0) = 1 \\ h(i) = \sum_{j=0}^{\infty} p(i, j) h(j) \end{cases}$$

Birth and death processes

$$(*) \quad \begin{cases} h(i) = \sum_{j \geq 0} p(i,j) h(j) \\ h(0) = 1 \end{cases}$$

$$p(i,j) = \begin{cases} p_i, & j = i+1 \\ q_i, & j = i-1 \\ 0, & \text{otherwise} \end{cases}$$

$$(*) \quad \begin{cases} h(0) = 1 \\ h(1) = p_1 h(2) + q_1 h(0) \\ h(2) = p_2 h(3) + q_2 h(1) \\ \vdots \\ h(i) = p_i h(i+1) + q_i h(i-1) \\ \vdots \end{cases}$$

$$\begin{cases} u(1) = h(1) - h(0) = h(1) - 1 \\ u(2) = \frac{q_1}{p_1} u(1) \\ u(3) = \frac{q_2}{p_2} u(2) = \frac{q_2 q_1}{p_2 p_1} u(1) \\ \vdots \\ u(i+1) = \frac{q_i}{p_i} u(i) = \frac{q_i \cdots q_1}{p_i \cdots p_1} u(1) \\ \vdots \end{cases}$$

$$p_i (h(i) - h(i+1)) = q_i (h(i-1) - h(i))$$

$\overset{\text{``}}{u(i+1)}$ $\overset{\text{``}}{u(i)}$

$$u(i+1) = \frac{q_i}{p_i} u_i$$

Denote

$$p_i := \frac{q_i \cdots q_1}{p_i \cdots p_1}$$

Birth and death processes

$$\left\{ \begin{array}{l} u(1) = u(1) \\ u(2) = p_1 u(1) \\ u(3) = p_2 u(1) \\ \vdots \\ u(i+1) = p_i u(1) \end{array} \right. \quad u(i) = h(i-1) - h(i)$$

Take the sum of the first i equations

$$h(0) - h(i) = (1 + p_1 + p_2 + \dots + p_{i-1}) u(1)$$

By Thm. 7.0 we need the minimal solution to $(*)$

Notice that $u(1)$ uniquely determines all $h(i)$

$$h(i) = h(0) - (1 + p_1 + p_2 + \dots + p_i) u(1)$$

and the minimal solution corresponds to maximal $u(1)$

- If $1 + \sum_{i=1}^{\infty} p_i = \infty$, then $u(1) = 0$ (otherwise $h(0) - h(1) > 1$)

In this case $h(0) - h(i) = 0 \quad \forall i \Rightarrow h(i) = 1$ for all i

no chance of survival

Birth and death processes

- If $1 + \sum_{i=1}^{\infty} p_i < \infty$, then for any $a \in [0, \frac{1}{1 + \sum_{i=1}^{\infty} p_i}]$

We get a solution to (*) by taking

$$\forall i \quad h(0) - h(i) = (1 + p_1 + p_2 + \dots + p_{i-1}) a$$

If $u(1) > \frac{1}{1 + \sum_{i=1}^{\infty} p_i}$, then for some m large enough

$$1 < (1 + \sum_{i=1}^m p_i) u(1) = h(0) - h(m) \leq 1$$

Therefore, $u(1) = \frac{1}{1 + \sum_{i=1}^{\infty} p_i}$ is the maximal allowable

value of $u(1)$, and the corresponding minimal

solution is $h(j) = 1 - (1 + \sum_{i=1}^{j-1} p_i) / (1 + \sum_{i=1}^{\infty} p_i) = \sum_{i=j}^{\infty} p_i / \sum_{i=1}^{\infty} p_i$ ■